

# Exploring the role of the halo-mass function for inferring astrophysical parameters during reionization

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Accepted 2024 August 12. Received 2024 July 16; in original form 2024 March 20

## ABSTRACT

Detecting the 21-cm signal at  $z \gtrsim 6$  will reveal insights into the properties of the first galaxies responsible for driving reionization. To extract this information, we perform parameter inference with three-dimensional simulations of the 21-cm signal embedded within a Bayesian inference pipeline. Presently, when performing inference, we must choose which sources of uncertainty to sample and which to hold fixed. Since the astrophysics of galaxies is much more uncertain than that of the underlying halo-mass function (HMF), we typically parametrize and model the former while fixing the latter. However, doing so may bias our inference of the galaxy properties. In this work, we explore the consequences of assuming an incorrect HMF and quantify the relative biases on our inferred astrophysical model parameters when considering the wrong HMF. We then relax this assumption by constructing a generalized five parameter HMF model and simultaneously recover it with our underlying astrophysical model. For this, we use 21CMFAST and perform simulation-based inference using marginal neural ratio estimation to learn the likelihood-to-evidence ratio with SWYFT. Using a mock 1000-h observation of the 21-cm power spectrum from the forthcoming Square Kilometre Array, conservatively assuming foreground wedge avoidance, we find that assuming the incorrect HMF can bias the recovered astrophysical parameters by up to  $\sim 3\text{--}4\sigma$  even when including independent information from observed luminosity functions. Using our generalized HMF model, although we recover our astrophysical parameters with a factor of  $\sim 2\text{--}4$  larger marginalized uncertainties, the constraints are unbiased, agnostic to the underlying HMF and therefore more conservative.

**Key words:** galaxies: high-redshift – intergalactic medium – diffuse radiation – early Universe – dark ages, reionization, first stars – cosmology: theory.

## 1 INTRODUCTION

The formation of the first stars and galaxies, referred to as the cosmic dawn (CD), signifies the end of the cosmic dark ages. The intergalactic medium (IGM), enshrouded by a neutral hydrogen fog following recombination renders the early Universe opaque to most forms of radiation. These first stars and galaxies, emitting copious amounts of ultraviolet (UV) photons, gradually eat away at this fog, leading to a complex morphology of ionized regions embedded in a neutral medium. Over time, as galaxies grow and become more abundant their combined UV photon output almost completely removes this neutral hydrogen fog. This process is referred to as the Epoch of Reionization (EoR).

Peering through this fog to directly explore the first generation of stars and galaxies is extremely difficult. Thankfully, we can indirectly infer the properties of the entire population of these first sources via their role in driving the EoR. This is achieved by measuring the temperature and spatial distribution of the neutral hydrogen in the IGM as it is first heated and then gradually disappears due to ionizations over cosmic time. We detect the neutral hydrogen via its 21-cm hyperfine spin-flip transition, which is measured relative to a background radiation source, such as the cosmic microwave background (CMB; see e.g. Gnedin & Ostriker 1997; Madau, Meiksin & Rees 1997; Shaver et al. 1999; Tozzi et al. 2000; Gnedin & Shaver 2004; Furlanetto, Oh & Briggs 2006; Morales & Wyithe 2010; Pritchard & Loeb 2012). Observing this 21-cm signal as a function of time (frequency) reveals a three-dimensional (3D), time-evolving picture of the ionization and thermal state of the IGM during the early Universe.

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Maximizing the wealth of astrophysical information in the 21-cm signal requires measuring this spatially varying signal, only achievable using large-scale radio interferometers. In recent years, several of these experiments have begun to report upper limits on the 21-cm signal via the power spectrum (PS; Mertens et al. 2020; Trott et al. 2020; Abdurashidova et al. 2022b; HERA Collaboration 2023). In turn, these upper limits have been scrutinized to begin to rule out credible regions of parameter space describing the galaxies responsible for reionization (e.g. Ghara et al. 2020; Mondal et al. 2020; Ghara et al. 2021; Greig et al. 2021a,b; Abdurashidova et al. 2022a; HERA Collaboration 2023). As we approach the first ever detection of the 21-cm signal, it is imperative that our theoretical frameworks and model assumptions are well characterized and understood.

For example, presently when inferring the properties of the galaxies that drive reionization we typically assume a fixed halo-mass function (HMF).<sup>1</sup> This choice is motivated by the fact that the uncertainties in the astrophysical properties describing the galaxies responsible for reionization are much larger than the uncertainties in the underlying HMF. Consequently, by fixing the HMF we can then minimize the computational burden of our inference pipelines. However, the cosmic 21-cm signal is extremely sensitive to the choice of HMF as demonstrated by Lopez-Honorez et al. (2016) for statistics such as the 21-cm global signal and 21-cm PS. In fact, the dependence of the HMF on the 21-cm PS can be recovered through the analytic framework of the halo model approach (Schneider, Giri & Mirocha 2021; Giri & Schneider 2022; Schneider, Schaeffer & Giri 2023). Therefore, given in practise we do not know the true HMF describing the Universe, by assuming a fixed HMF within our inference pipelines our inferred insights into the astrophysical properties describing these galaxies may be biased, especially given the sensitivity of the cosmic 21-cm signal. In this work, we will explore how biased our astrophysical inferences can be due to this assumption of a fixed HMF by varying several theoretical models with incorrectly assumed HMFs relative to a fiducial mock observation with known HMF.

Mirocha, Lamarre & Liu (2021) explored the role of several of these theoretical modelling uncertainties, including the impact of the HMF, on our ability to infer the galaxy properties using the 21-cm global signal. Along with the stellar population synthesis model, the choice of assumed HMF was observed to have a significant impact on the inferred astrophysics. However, the numerical modelling of the 21-cm signal in Mirocha et al. (2021) employed a simple two-zone IGM model and forgoes modelling the complex 3D nature of the cosmic signal. Therefore, we revisit this, exploring the impact of assuming a fixed HMF in the context of observations of the 21-cm PS, which requires modelling the full 3D signal.

Traditionally, to tackle this Bayesian inference problem we would employ a tool such as 21CMC (Greig & Mesinger 2015, 2017; Greig & Mesinger 2018; Park et al. 2019) which performs on-the-fly 3D seminumerical reionization simulations using 21CMFAST (Mesinger & Furlanetto 2007; Mesinger, Furlanetto & Cen 2011; Murray et al. 2020) within a Monte Carlo Markov Chain (MCMC) framework. However, such an approach can be computationally intensive, especially as the number of free parameters increase or in this instance where we are interested in varying the HMF requiring a new MCMC per HMF. Therefore, in this work we

utilize likelihood-free or simulation-based inference (SBI; see e.g. Cranmer, Brehmer & Louppe 2020 for a recent review) that adopts machine learning principles to either bypass or model the general likelihood function (avoiding the typically Gaussian assumptions to be able to derive an analytic functional form of the likelihood; see e.g. Prelogović & Mesinger 2023 for detailed discussions). SBI directly connects the forward modelled simulated data to the posterior distribution of the input model parameters or to first ‘learn’ the likelihood from the same forward modelled data which can then be used within an MCMC framework to obtain the model posteriors. There are several benefits of applying these approaches including the removal of the need for an assumed analytic likelihood which relaxes many assumptions and in the process it enables complex (non-Gaussian) summary statistics to be more readily accessed in a Bayesian framework. Further, the forward modelled simulated data sets can be more efficient at sampling the prior volume (removing redundant model evaluations that occur in an MCMC), leading to fewer simulated models to achieve comparable or better performance. In recent years, these approaches have gained in popularity for tackling astrophysical inference from the cosmic 21-cm signal (e.g. Zhao et al. 2022a; Zhao, Mao & Wandelt 2022b; Prelogović & Mesinger 2023; Saxena et al. 2023).

In this work, we choose to follow the approach of Saxena et al. (2023) and perform our parameter inference using marginal neutral ratio estimation (MNRE; Miller et al. 2021) using the publicly available PYTHON package, SWYFT<sup>2</sup> (Miller et al. 2022). Rather than learning the likelihood, in MNRE the goal is instead to learn the marginal likelihood-to-evidence ratios for each individual parameter combination through its own unique neural network. This approach can be more efficient than sampling the full posterior using an MCMC (Saxena et al. 2023). Owing to these benefits, this approach has already been applied for several cosmological applications including the CMB (Cole et al. 2022), gravitational lensing (Anau Montel et al. 2023; Coogan et al. 2024), gravitational waves (Bhardwaj et al. 2023; Gagnon-Hartman, Ruan & Haggard 2023), Milky Way stellar streams (Alvey, Gerdes & Weniger 2023), and supernovae cosmology (Karchev, Trotta & Weniger 2023).

Specifically, for our work we shall utilize SWYFT to perform astrophysical parameter inference from a mock observation of the 21-cm PS from the forthcoming Square Kilometre Array (SKA; Mellema et al. 2013; Koopmans et al. 2015). For this, we will utilize 21CMFASTs flexible galaxy parametrization (Park et al. 2019) to make use of UV luminosity functions as additional observational priors. We shall perform several inferences, varying the underlying HMF in order to quantify the bias in the inferred parameters as a result of our assumptions. Following this, making use of SWYFTs efficiency, we then look to relax our assumption of a fixed HMF by jointly recovering our astrophysical model parameters and those of a generalized HMF functional form.

The remainder of this paper is organized as follows. In Section 2, we summarize our 21-cm simulations using 21CMFAST and provide summary statistics of the 21-cm signal in order to highlight its sensitivity to the underlying choice of HMF. In Section 3, we describe our approach for constructing our noisy, simulated data for performing SBI with SWYFT. In Section 4, we explore the impact of assuming a fixed HMF on the inferred astrophysical model parameters before considering a generalized function form for the HMF in Section 5 and simultaneously recovering our astrophysical and HMF model parameters. Finally, in Section 6 we provide our closing remarks.

<sup>1</sup>We also typically assume fixed cosmological parameters, which if allowed to vary will also have an impact on the inferred astrophysical parameters (Kern et al. 2017).

<sup>2</sup><https://github.com/undark-lab/swyft>

Unless stated otherwise, all quantities are in co-moving units and we adopt the cosmological parameters:  $(\Omega_\Lambda, \Omega_M, \Omega_b, n, \sigma_8, H_0) = (0.69, 0.31, 0.048, 0.97, 0.81, 68 \text{ km s}^{-1} \text{ Mpc}^{-1})$ , consistent with recent results from the Planck mission (Planck Collaboration VI 2020).

## 2 SIMULATING THE 21-CM SIGNAL

For simulating the 3D 21-cm signal during reionization we use the semi-numerical simulation code 21CMFAST<sup>3</sup> (Mesinger & Furlanetto 2007; Mesinger et al. 2011), specifically the latest public release, v3 (Murray et al. 2020). In particular, we adopt the flexible galaxy parametrization of Park et al. (2019) that describes the UV and X-ray properties of the galaxies responsible for reionization. Below, we outline the key concepts for simulating the 21-cm signal, with particular focus on the astrophysical parameters that we seek to recover in our parameter inference pipeline and their dependence on the HMF. For more detailed discussions, please refer to these aforementioned publications.

### 2.1 Galaxy UV properties

The stellar mass,  $M_*$ , of a galaxy is assumed to be directly proportional to its host halo mass,  $M_h$  (e.g. Kuhlen & Faucher-Giguère 2012; Dayal et al. 2014; Behroozi & Silk 2015; Mitra, Choudhury & Ferrara 2015; Mutch et al. 2016; Ocvirk et al. 2016; Sun & Furlanetto 2016; Yue, Ferrara & Xu 2016; Hutter et al. 2021),

$$M_*(M_h) = f_* \left( \frac{\Omega_b}{\Omega_m} \right) M_h, \quad (1)$$

with the fraction of galactic gas in stars,  $f_*$ , also dependent on the host halo mass via a power-law<sup>4</sup> characterized by two free model parameters, the index,  $\alpha_*$ , and the expression is normalized via  $f_{*,10}$  for a host dark matter halo mass of  $10^{10} M_\odot$ ,

$$f_* = f_{*,10} \left( \frac{M_h}{10^{10} M_\odot} \right)^{\alpha_*}. \quad (2)$$

Dividing this stellar mass by a characteristic time-scale,  $t_*$ , then yields the star formation rate (SFR) for these galaxies,

$$\dot{M}_*(M_h, z) = \frac{M_*}{t_* H^{-1}(z)}, \quad (3)$$

where  $H^{-1}(z)$  is the Hubble time and  $t_* \in [0.05, 1]$ .

The fraction of UV photons escaping into the IGM,  $f_{\text{esc}}$ , is also assumed to be governed by a power-law relation with halo mass,

$$f_{\text{esc}} = f_{\text{esc},10} \left( \frac{M_h}{10^{10} M_\odot} \right)^{\alpha_{\text{esc}}}, \quad (4)$$

yielding an additional two parameter, index,  $\alpha_{\text{esc}}$ , and normalization,  $f_{\text{esc},10}$ .

Importantly, not all haloes are capable of contributing to reionization owing to either internal feedback and/or inefficient cooling that can stymie star formation in low-mass haloes. To mimic this behaviour, an effective duty-cycle is adopted,

$$f_{\text{duty}} = \exp \left( -\frac{M_{\text{turn}}}{M_h} \right). \quad (5)$$

<sup>3</sup><https://github.com/21cmfast/21cmFAST>

<sup>4</sup>This follows directly from the mean behaviour of  $M_*$  and  $M_h$  obtained from both semi-empirical fits to observations (e.g. Harikane et al. 2016; Tacchella et al. 2018; Behroozi et al. 2019; Stefanon et al. 2021) and semi-analytic model predictions (e.g. Mutch et al. 2016; Yung et al. 2019; Hutter et al. 2021).

where  $(1 - f_{\text{duty}})$  corresponds to the fraction of suppressed star-forming galaxies below a characteristic mass scale  $M_{\text{turn}}$  (e.g. Giroux, Sutherland & Shull 1994; Shapiro, Giroux & Babul 1994; Hui & Gnedin 1997; Barkana & Loeb 2001; Springel & Hernquist 2003; Mesinger & Dijkstra 2008; Okamoto, Gao & Theuns 2008; Sobacchi & Mesinger 2013a,b).

The primary impact of our assumed HMF is through the production rate of UV photons, which directly feeds into the corresponding UV luminosity functions (LFs). The above galaxy parametrization connects to UV LFs by first computing the non-ionizing UV LFs:

$$\phi(M_{\text{UV}}) = \left[ f_{\text{duty}} \frac{dn}{dM_h} \right] \left| \frac{dM_h}{dM_{\text{UV}}} \right|. \quad (6)$$

The quantity in square brackets corresponds to the number density of active star-forming galaxies and is directly proportional to the HMF. Therefore, including the UV LFs when exploring the role of assuming a fixed HMF will be important. The final term performs the conversion between halo mass and UV magnitude. Our defined SFR (equation 3) is then related to the UV luminosity ( $L_{\text{UV}}$ ) via

$$\dot{M}_*(M_h, z) = \mathcal{K}_{\text{UV}} \times L_{\text{UV}}, \quad (7)$$

where  $\mathcal{K}_{\text{UV}} = 1.15 \times 10^{-28} M_\odot \text{ yr}^{-1} / \text{erg s}^{-1} \text{ Hz}^{-1}$  is a conversion factor (Madau & Dickinson 2014) and the UV luminosity is related to AB magnitude via the standard relation (Oke & Gunn 1983):

$$\log_{10} \left( \frac{L_{\text{UV}}}{\text{erg s}^{-1} \text{ Hz}^{-1}} \right) = 0.4 \times (51.63 - M_{\text{UV}}). \quad (8)$$

In summary, this model contains six free parameters to describe the UV properties of the galaxies, namely,  $f_{*,10}$ ,  $\alpha_*$ ,  $f_{\text{esc},10}$ ,  $\alpha_{\text{esc}}$ ,  $\alpha_{\text{ast}}$  and  $M_{\text{turn}}$ .

### 2.2 Galaxy X-ray properties

In addition to UV photons, these first galaxies emit X-ray photons originating from stellar remnants from earlier episodes of star formation. These escaping X-ray photons, with their long mean-free paths, are capable of traversing large distances and heating the adiabatically cooling IGM. Within 21CMFAST, X-ray heating is modelled by first computing the angle-averaged specific X-ray intensity,  $J(\mathbf{x}, E, z)$ , (in  $\text{erg s}^{-1} \text{ keV}^{-1} \text{ cm}^{-2} \text{ sr}^{-1}$ ) within each simulation voxel,

$$J(\mathbf{x}, E, z) = \frac{(1+z)^3}{4\pi} \int_z^\infty dz' \frac{cdt}{dz'} \epsilon_X e^{-\tau}. \quad (9)$$

This calculates the contribution from the co-moving X-ray specific emissivity,  $\epsilon_X(\mathbf{x}, E_e, z')$ , emitted at earlier times,  $E_e = E(1+z')/(1+z)$ , while also accounting for their attenuation by the IGM,  $e^{-\tau}$ , as they traverse the vast distances through the Universe. The specific emitted emissivity is then

$$\epsilon_X(\mathbf{x}, E_e, z') = \frac{L_X}{\text{SFR}} \left[ (1 + \bar{\delta}_{\text{nl}}) \int_0^\infty dM_h \frac{dn}{dM_h} f_{\text{duty}} \dot{M}_* \right], \quad (10)$$

where  $\bar{\delta}_{\text{nl}}$  is the mean, non-linear overdensity in a shell centred on the simulation cell  $(\mathbf{x}, z)$ ,  $L_X/\text{SFR}$  ( $\text{erg s}^{-1} \text{ keV}^{-1} M_\odot^{-1} \text{ yr}$ ) is the specific X-ray luminosity per unit star formation escaping the host galaxies and the quantity in square brackets is the star formation rate density (SFRD) along the light-cone. This latter quantity depends on our assumed choice for the HMF highlighting that our modelled X-ray emissivity is directly impacted by our HMF choice. Either a change to the HMF could be compensated for by a corresponding change to  $L_X/\text{SFR}$  or through the stellar properties of the UV galaxies, namely  $M_{\text{turn}}$  (through  $f_{\text{duty}}$ ),  $f_{*,10}$ ,  $\alpha_*$ , or  $t_*$ .

$L_X/\text{SFR}$  acts as the normalization of the X-ray emissivity and is modelled with a power-law dependence with photon energy,  $L_X \propto E^{-\alpha_X}$ , with  $\alpha_X$  denoting the index of the spectral energy distribution describing the source of X-rays. In this work, high-mass X-ray binaries are assumed to be dominant, thus we adopt  $\alpha_X = 1$  consistent with local Universe observations (Mineo, Gilfanov & Sunyaev 2012; Fragos et al. 2013; Pacucci et al. 2014). Finally, this specific X-ray luminosity is then normalized by the integrated soft-band ( $< 2$  keV) luminosity per SFR (in  $\text{erg s}^{-1}\text{M}_\odot^{-1}$  yr),

$$L_{X<2\text{keV}}/\text{SFR} = \int_{E_0}^{2\text{keV}} dE_e L_X/\text{SFR}, \quad (11)$$

where  $E_0$  corresponds to the energy threshold below which X-ray photons are absorbed by the host galaxy. Both  $L_{X<2\text{keV}}/\text{SFR}$  and  $E_0$  are free parameters in our astrophysical model.

### 2.3 Ionization and thermal state of the IGM

Computing the thermal and ionization states of the IGM requires the velocity and evolved density fields which are determined using second-order Lagrange perturbation theory (Scoccimarro 1998). First, the thermal state of the neutral IGM is calculated by determining the IGM spin temperature,  $T_S$ , which is obtained by self-consistently computing X-ray heating and ionizations along with heating and cooling due to structure formation, Compton scattering by CMB photons and heating due to partial ionizations.  $T_S$  is set by the weighted mean,

$$T_S^{-1} = \frac{T_{\text{CMB}}^{-1} + x_\alpha T_\alpha^{-1} + x_c T_K^{-1}}{1 + x_\alpha + x_c}, \quad (12)$$

where  $T_K$ ,  $T_\alpha$ , and  $T_{\text{CMB}}$  are the gas, Lyman  $\alpha$  ( $\text{Ly}\alpha$ ) colour, and CMB temperatures, respectively, and  $T_S$  depends both on the local gas density and the  $\text{Ly}\alpha$  radiation intensity. The  $\text{Ly}\alpha$  background is the cumulative sum of X-ray excitations of neutral hydrogen atoms and the direct stellar emission of Lyman band photons by the first sources with  $x_\alpha$  denoting the Wouthuysen-Field coupling coefficient (Wouthuysen 1952; Field 1958) and  $x_c$  the collisional coupling coefficient between the free electrons and CMB photons.

The ionization field is then determined by applying excursion-set theory (Furlanetto, Zaldarriaga & Hernquist 2004) on the evolved density field which compares the cumulative number of ionizing photons,  $n_{\text{ion}}$ , to the total number of neutral hydrogen atoms plus cumulative recombinations,  $\bar{n}_{\text{rec}}$  (Sobacchi & Mesinger 2014) within spheres of decreasing radii,  $R$ , and corresponding overdensity,  $\delta_R$ . An individual simulation voxel is then considered ionized if

$$\bar{n}_{\text{ion}}(\mathbf{x}, z|R, \delta_R) \geq (1 + \bar{n}_{\text{rec}})(1 - \bar{x}_e), \quad (13)$$

where  $(1 - \bar{x}_e)$  accounts for secondary ionizations by X-ray photons. The cumulative number of ionizations is determined by

$$n_{\text{ion}} = \bar{\rho}_b^{-1} \int_0^\infty dM_h \frac{dn(M_h, z|R, \delta_R)}{dM_h} f_{\text{duty}} \dot{M}_* f_{\text{esc}} N_{\gamma/b}, \quad (14)$$

where  $\bar{\rho}_b$  is the mean baryon density and the total number of ionizing photons per stellar baryon is given by  $N_{\gamma/b}$  which is assumed to be  $N_{\gamma/b} = 5000$  consistent with a Salpeter initial mass function (Salpeter 1955). As one would expect,  $n_{\text{ion}}$  is directly connected to our adopted HMF. Therefore, the relative amplitude and slope of the HMF (i.e. number density of haloes of a given mass) will have a notable impact on the total number of ionizations and thus primarily the ionization state of the IGM. Consequently, this implies any of the six free model parameters describing the UV galaxies could be impacted by the assumed choice of HMF.

### 2.4 21-cm brightness temperature

The intensity of the observed 21-cm radiation, or brightness temperature,  $\delta T_b(\nu)$ , is set by the optical depth,  $\tau_{\nu_0}$ , of the neutral hydrogen gas (where  $\nu_0$  is the frequency of the 21-cm signal) and the background radiation passing through it (e.g. CMB; Furlanetto et al. 2006);

$$\delta T_b(\nu) = \frac{T_S - T_{\text{CMB}}(z)}{1+z} (1 - e^{-\tau_{\nu_0}}) \text{ mK}, \quad (15)$$

and

$$\tau_{\nu_0} \propto (1 + \delta_{\text{nl}})(1+z)^{3/2} \frac{x_{\text{H1}}}{T_S} \left( \frac{H}{dv_r/dr + H} \right). \quad (16)$$

The optical depth depends on the neutral hydrogen fraction,  $x_{\text{H1}}$ , local gas overdensity,  $\delta_{\text{nl}} \equiv \rho/\bar{\rho} - 1$ , the Hubble parameter,  $H(z)$ , and the line-of-sight gradient of the peculiar velocity. This quantity is evaluated at the redshift  $z = \nu_0/\nu - 1$  and for ease of notation we have dropped the explicit spatial dependence of the quantities.

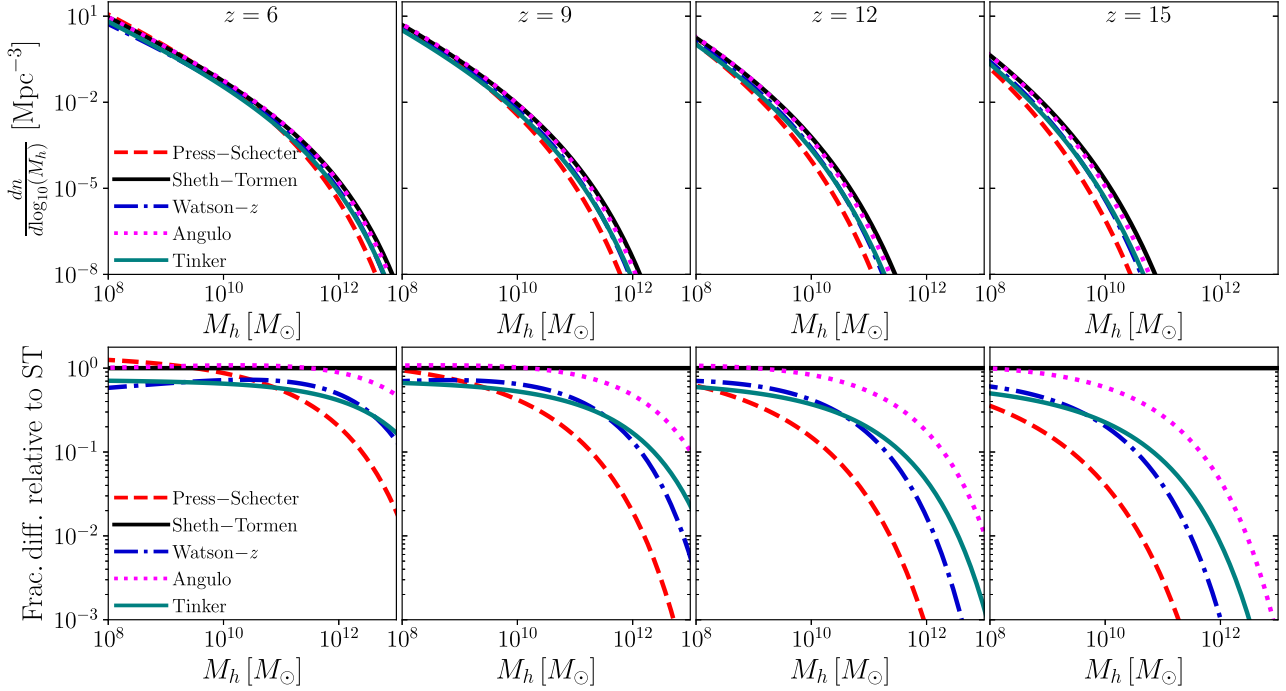
### 2.5 The halo-mass function

As outlined previously, the timing and bias of the sources producing UV and X-ray photons is degenerate with our modelled astrophysical parameters which is in turn degenerate with the HMF. In 21CMFAST, the HMF is explicitly calculated using an analytic expression; therefore, it is straightforward to interchange for various analytic HMFs in the literature. By default, 21CMFAST assumes the Sheth, Mo & Tormen (2001, hereafter ST) ellipsoidal collapse model to compute the number density of haloes. Throughout this work, we adopt ST as our default HMF. For completeness of the purely analytically derived HMFs we shall also consider the Press & Schechter (1974, hereafter PS) HMF.

Following these, within the literature there have been many analytic fitting functions created to match outputs of  $N$ -body simulations (e.g. Jenkins et al. 2001; Warren et al. 2006; Reed et al. 2007; Tinker et al. 2008; Tinker et al. 2010; Angulo et al. 2012; Watson et al. 2013; Diemer 2020). These can vary considerably in their functional forms with differences arising from the fundamentals of the  $N$ -body simulations (e.g. gravity solver and/or definition of what constitutes a halo along with the halo finder itself). Importantly, the vast majority of these  $N$ -body simulations are not designed for high- $z$  studies, which is not unsurprising given the wealth of observations at low- $z$  (and lack thereof at high- $z$ ) from which to calibrate the  $N$ -body results against.<sup>5</sup> Therefore, the extrapolation of these HMFs out to the high- $z$ 's necessary for studying reionization can lead to significant differences in the number density of sources that can have a significant impact on the inferred astrophysical parameters (Mirocha et al. 2021).

In addition to the ST and PS HMFs, we also consider the analytic fitting functions of Tinker et al. (2010, hereafter Tinker), Angulo et al. (2012, hereafter Angulo) and the redshift evolving Watson et al. (2013, hereafter Watson- $z$ ). The choice is somewhat arbitrary, but is designed to broadly cover the breadth of variation in the 21-cm signal during reionization. For example, we found the Diemer (2020) model to be very similar to that of Watson- $z$  while equally the Angulo and Jenkins et al. (2001) models were also very similar.

<sup>5</sup>In fact, even at low redshifts the HMFs are known to differ non-negligibly (Murray, Power & Robotham 2013).



**Figure 1.** The HMFs at  $z = 6, 9, 12$  and  $15$  for the five different models we consider in this work. In particular, we consider the Sheth–Tormen (black solid), Press–Schechter (red dashed), redshift evolving Watson (blue dot–dashed), Angulo (magenta dotted), and the Tinker (teal triple-dot–dashed) HMFs. In the bottom row, we provide the fractional difference between the various HMFs and that of the ST HMF.

**Table 1.** The astrophysical model parameters for our mock 21-cm observation, assuming a Sheth et al. (2001) HMF, based on the constraints from Qin et al. (2021b).

	$\log_{10}(f_{*,10})$	$\alpha_*$	$\log_{10}(f_{\text{esc},10})$	$\alpha_{\text{esc}}$	$\log_{10}(M_{\text{turn}})$ ( $M_{\odot}$ )	$t_*$	$\log_{10} \frac{L_{X < 2 \text{ keV}}}{\text{SFR}}$ ( $\text{erg s}^{-1} M_{\odot}^{-1} \text{yr}^{-1}$ )	$E_0$ (keV)
Mock Observation	−1.1	0.5	−1.30	−0.35	8.55	0.5	40.5	0.5
Prior ranges	[−3.0, 0.0]	[−0.5, 1.0]	[−3.0, 0.0]	[−1.0, 0.5]	[8.0, 10.0]	[0.05, 1.0]	[38.0, 42.0]	[0.1, 1.5]

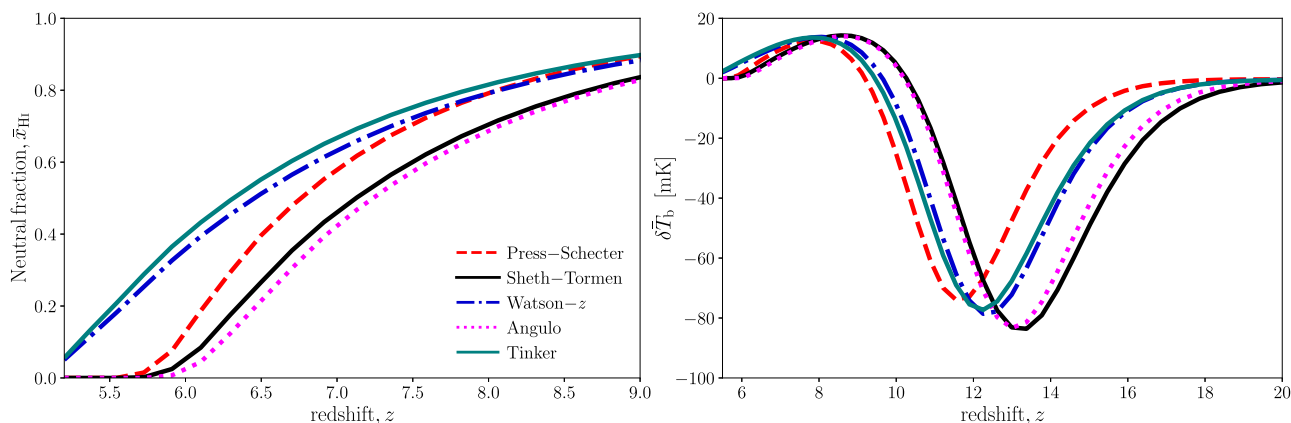
In Fig. 1, we provide the corresponding HMFs for our selected models at  $z = 6, 9, 12$ , and  $15$ . Additionally, for ease of comparison we also provide the fractional difference between the various HMFs and the ST HMF. At  $z = 6$ , all models are fairly consistent with one another up until halo masses of  $10^{11} M_{\odot}$  following which the various models begin to diverge for increasing mass. Although these large mass haloes do not significantly contribute to the total number of ionizing photons, they contribute to setting  $f_*$  through the UV LFs and thus small differences in the HMF on these mass scales can still bleed into our inference of the astrophysical parameters. As noted earlier, as we extrapolate these models out to higher redshifts, the relative differences are amplified and extend throughout the entire mass range. At  $z = 9$ , the HMFs are only consistent below  $10^{9.5} M_{\odot}$ , whereas by  $z = 15$  significant differences occur over the entire mass range. Further, the slope of the different models also differs, resulting in different abundances of sources within different mass bins. Combined, these differences will have a sizeable impact on the 21-cm signal.

## 2.6 Mock 21-cm observation

In order to quantify the relative impact of our HMF choice on our inferred astrophysical parameters, we must define a mock

observation of the 21-cm signal.<sup>6</sup> For this, we adopt a set of astrophysical parameters that closely resemble the recovered model of Qin et al. (2021b) based on the cumulative distribution functions of the Ly $\alpha$  forest (Bosman et al. 2018) combined with existing constraints on the electron scattering optical depth,  $\tau_e$  from the CMB (Planck Collaboration VI 2020), UV LFs (see Section 2.6.3) and dark fraction (McGreer, Mesinger & D’Odorico 2015). In Table 1, we summarize the corresponding model parameters of our mock observation, along with the adopted prior ranges for each of these parameters based on previous work (Park et al. 2019; Greig et al. 2021b). Note, throughout this work, these astrophysical parameters always remained fixed, and we only ever change the underlying HMF when differentiating models. For this mock observation we adopt a ST HMF to describe the number density of ionizing sources. Note,

<sup>6</sup>Theoretically one could also consider the impact of varying the underlying transfer function used for modelling the matter power spectrum. For example, using 21CMFAST we assume a matter power spectrum defined by (Eisenstein & Hu 1999). However, more accurate transfer functions exist which rely on Boltzmann solvers such as CAMB (Lewis, Challinor & Lasenby 2000) or CLASS (Blas, Lesgourgues & Tram 2011). Relatively speaking the differences are more minor, with differences consistent with those between the ST and Angulo HMFs. Nevertheless, they will contribute marginally to biasing our inferred results.



**Figure 2.** The global history of reionization (left panel) and the mean brightness temperature (right panel) for the same astrophysical model parameters (see Table 1) but varying the underlying HMF. In particular, we consider the Sheth–Tormen (black solid; our mock 21-cm observation), Press–Schechter (red dashed), redshift evolving Watson (blue dot–dashed), Angulo (magenta dotted), and the Tinker (teal triple–dot–dashed) HMFs. Note, for this visual comparison, all 21-cm light-cones are generated using the same initial conditions for our fiducial 250 Mpc simulations.

for this section, to ease qualitative discussions, all 21-cm models are generated assuming the same initial conditions.

### 2.6.1 Global signal

In Fig. 2, we demonstrate the global history of reionization (left panel) and the mean brightness temperature signal (right panel) as a function of redshift for our mock 21-cm model assuming a ST HMF (solid black curves). Additionally, we also provide the corresponding curves for our four other HMFs; PS (red, dashed), Watson- $z$  (blue, dot–dashed), Angulo (magenta, dotted), and Tinker (teal, triple–dot–dashed). Immediately evident from this figure are the considerably large differences in the corresponding timing of reionization along with both the amplitude of the global 21-cm signal and the corresponding redshifts of the various features.

Relative to our fiducial model (ST), the mid-point of reionization can differ by as much as  $\Delta z \sim 1$  depending on the choice of HMF, with both the Tinker and Watson- $z$  models resulting in a delayed and slower reionization. This is due to the  $\sim 50$  per cent reduction in the number of sources (e.g. Fig. 1) at the low-mass end. The differences in the HMF choices correspond to a shift in the electron scattering optical depth of at most  $\Delta \tau_e \pm 0.005$  relative to the fiducial model of  $\tau_e \sim 0.05$ . Equally, this difference corresponds to a  $\Delta z \sim 1$  shift in the features in the global brightness temperature. The PS model exhibits a notably different reionization history and 21-cm global signal. At high- $z$ , PS considerably under predicts the number of haloes, resulting in the most delayed Ly $\alpha$  decoupling ( $z \sim 15$ ) and absorption trough ( $z \sim 11$ ). However, the PS HMF more rapidly evolves on the low-mass end with decreasing redshift, therefore although reionization starts latest for PS, its duration occurs on the shortest time-scale for any model due to the more rapidly increasing abundance of sources (largest gradient in the reionization history). It is quite clear from these observations that the various HMFs can produce relatively drastic differences in the globally averaged 21-cm signal.

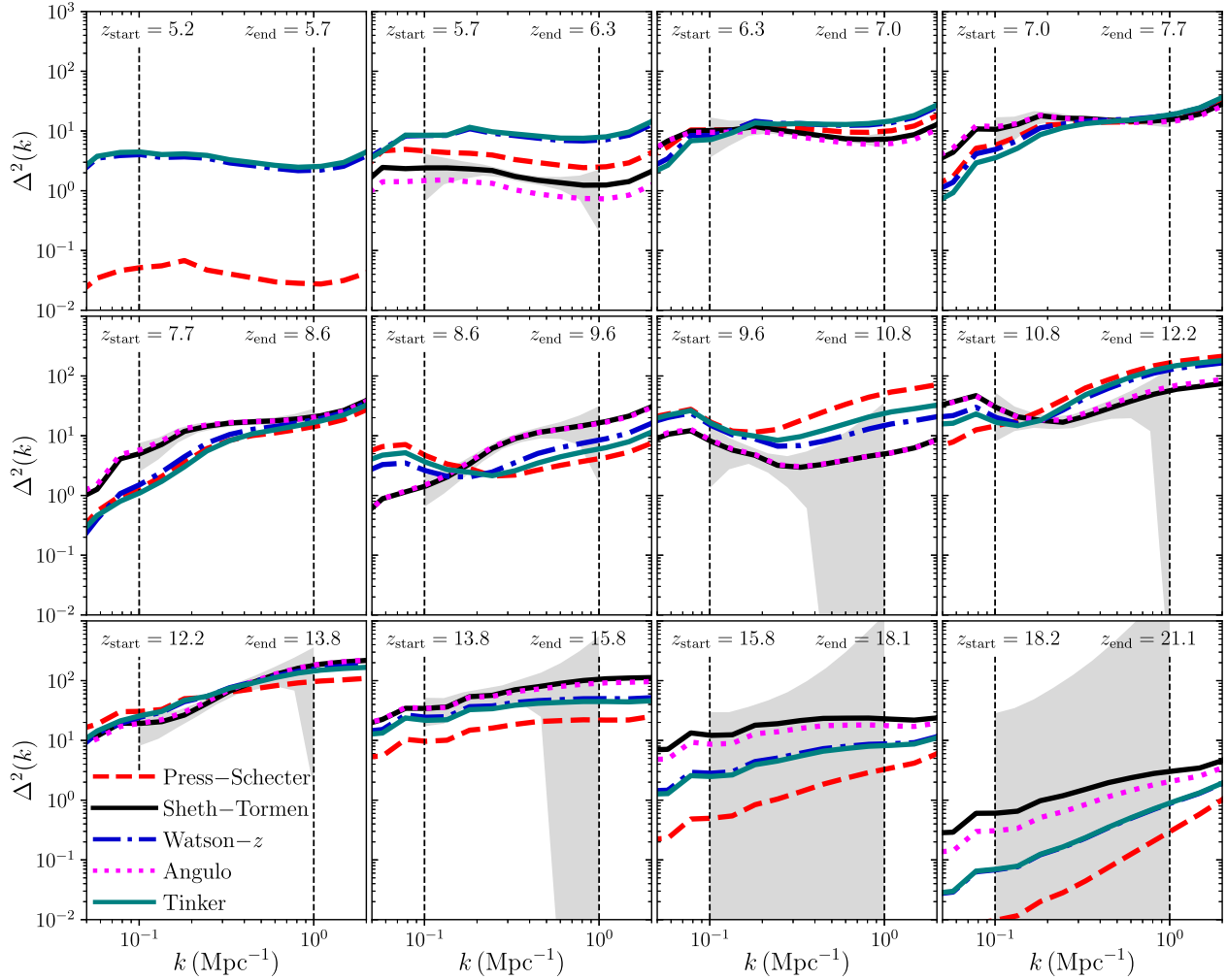
### 2.6.2 21-cm power spectrum

In Fig. 3, we provide the 21-cm PS for our astrophysical model, again varying the underlying HMF. For this, we take the 3D 21-cm

light-cone output by 21CMFAST and break it up into pieces along the line of sight, with the start and end redshifts of each cube denoted at the top of each panel. The vertical dashed lines correspond to the spatial scales between which we use the 21-cm PS for our astrophysical parameter inference ( $k = 0.1$  and  $k = 1.0$  Mpc $^{-1}$ ). Finally, the shaded regions denote the 68th percentile uncertainties obtained from a 1000 h observation with the SKA using 21CMSense (Poher et al. 2013, 2014). These are provided here purely as a visual aid to provide context to the relative differences in the 21-cm PS due to the choice of HMF. Further, for this figure, we assume perfect foreground removal in order to highlight variations to the 21-cm PS coming directly from the underlying changes to the HMF. In Section 3.2, we outline how we model the instrumental effects and include them in our forward-modelled simulations for our parameter inference.

Like previously, for the same fixed astrophysical parameter set the choice of HMF can have a significant impact on the 21-cm PS. Because the choice of HMF impacts the relative timing of the various 21-cm milestones (e.g. reionization, X-ray heating or Ly $\alpha$  decoupling), for the same fixed redshift (frequency) we are effectively comparing the 21-cm signal at a different evolutionary stage. For example, in most panels the PS, Watson- $z$  and Tinker models lag behind the mock ST model by almost one entire panel. Relative to the observational uncertainties from the SKA, this can result in differences in excess of several  $\sigma$ , especially during reionization (e.g.  $z \lesssim 10$ ), when the observational uncertainties are smallest. Even at higher redshifts, when the observational uncertainties are much larger due to increasing thermal noise, certain models can still differ by several  $\sigma$  due to the relatively large amplitude signal during the X-ray heating epoch. Therefore, when performing inference and changing the underlying HMF we will be biased to astrophysical parameters that can result in the timing of these cosmic milestones occurring earlier to match that of our fiducial ST model.

Even the Angulo model, which is very similar to the fiducial ST model can differ in excess of the SKA observational uncertainties. For example, during reionization (e.g.  $z \lesssim 10$ ) the Angulo HMF produces slightly more haloes relative to the ST HMF at  $\lesssim 10^{10} M_{\odot}$  (see e.g. Fig. 1). As a result, reionization begins slightly earlier and proceeds slightly faster (see Fig. 2). Even by  $z < 7$  these marginal differences in the HMF can produce 21-cm PS that differ by more



**Figure 3.** The 21-cm power spectrum (PS) obtained from a 3D 21-cm light-cone generated by 21CMFAST for our fiducial astrophysical parameter set, but varying the underlying HMF. Here, we consider the 21-cm PS obtained from equal comoving length chunks along the 21-cm light-cone. Note, for this we show the 21-cm PS obtained from the full simulation cube, not including any instrumental effect (e.g. foreground wedge) or thermal noise. However, as a visual guide the shaded regions correspond to the 68th percentile uncertainties for a mock 1000 h observation with the SKA. The vertical dashed lines denote the Fourier scales used within our inference pipeline ( $k = 0.1$  and  $k = 1.0 \text{ Mpc}^{-1}$ , respectively). Note, for this visual comparison, all 21-cm light-cones are generated using the same initial conditions.

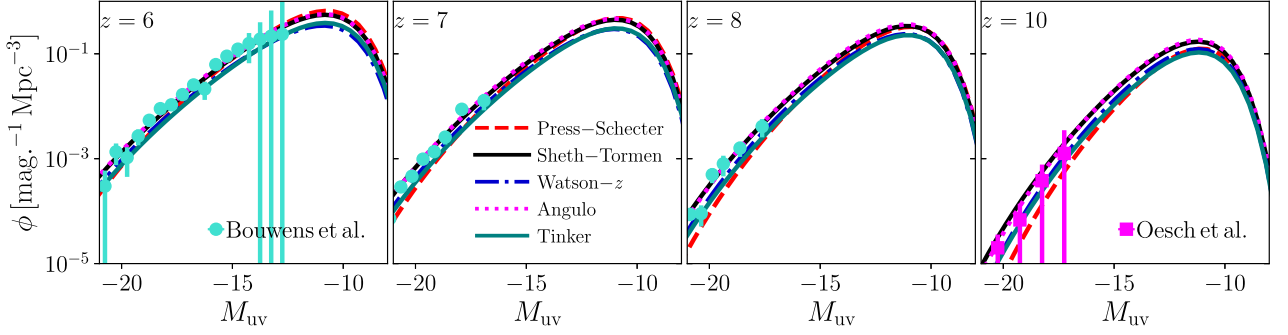
than the observational uncertainties. Therefore, the Angulo model will serve as a demonstration of how sensitive our HMF choice is to the inferred astrophysical parameters.

### 2.6.3 UV LFs

Finally, we also consider the impact of our choice of HMF on the UV LFs. In Park et al. (2019), it was demonstrated that the unique and complimentary information provided by observed high- $z$  UV LFs can improve on the overall constraining power of the underlying model (see also Mirocha, Furlanetto & Sun 2017). Essentially, this improvement stems from the sensitivity of the UV LFs to the stellar parameters,  $f_*$ , breaking the degeneracy with the escape-fraction,  $f_{\text{esc}}$  when it comes to the UV ionizations. Importantly, since the UV LFs also depend on the HMF, our relative improvements also depend on our assumed HMF.

In Fig. 4, we demonstrate the sensitivity of the modelled UV LFs obtained using 21CMFAST for our mock astrophysical parameter set varying only our choice of HMF. For reference, we additionally

include a selection of observed UV LFs at  $z = 6$  (Bouwens et al. 2017),  $z = 7$  and  $8$  (Bouwens et al. 2015) and  $z = 10$  (Oesch et al. 2018). Note, this choice follows that of Park et al. (2019), based on limiting the systematic differences between the various results in the literature due to how the statistical and observational uncertainties are determined. Ideally, one should combine all results in the literature to obtain a mean UV LF that characterizes the scatter across the various observations and pipelines, resulting in a more conservative and unbiased UV LF. However, for now we restrict our results to these existing UV LFs. Additionally, for the purposes of our modelling we ignore the impact of dust on the bright UV galaxies, imposing a limit of  $M_{\text{UV}} < -20$  below which we argue the galaxies are dust-free (see Park et al. 2019, for further discussions). However, this may not be the case, and any non-zero dust contribution would result in a reddening of the UV slope biasing the inference on  $f_*$ . Therefore, in future we aim to explore the inclusion of dust in our model and to make use of the bright-end of the UV LFs. For example, one could envisage utilizing either empirical or physically-motivated prescriptions for dust and their connection to the UV properties of



**Figure 4.** The corresponding UV LFs, assuming the same fixed astrophysical parameter set, but varying the underlying HMF. We show the UV LFs at  $z = 6, 7, 8$  and  $10$  which correspond to the redshifts of existing observational data.

the galaxies that have been extensively explored in the literature (e.g. Meurer, Heckman & Calzetti 1999; Wilkins et al. 2013; Bouwens et al. 2014; Khakhaleva-Li & Gnedin 2016; Mancini et al. 2016; Ma et al. 2019; Qiu et al. 2019; Mirocha, Mason & Stark 2020; Zhao & Furlanetto 2024).

For these specific UV LFs, we can see that the variations due to the HMFs are larger than the corresponding statistical uncertainties, therefore the UV LFs have sufficient constraining power to differentiate between the HMF models. Note, however, this would be notably reduced if we averaged across all UV LFs and the various systematic uncertainties. For reference, the turn-over in the UV LFs corresponds to our choice for  $M_{\text{turn}} = 5 \times 10^8 M_{\odot}$ , indicating that the UV LFs are sensitive primarily to the HMF on mass scales greater than this. As highlighted earlier, the extrapolation of the HMFs out to higher redshifts increases the relative differences, which we observe in the case of the UV LFs. Finally, as noted previously changes to the underlying HMF can be compensated by increasing the number of UV photons escaping the galaxies. However, since the UV LFs constrain the stellar contribution through  $f_*$ , this implies that when we combine observations of the UV LFs with the 21-cm signal, the largest biases should appear via the escape fraction,  $f_{\text{esc}}$  as the UV LFs are independent of  $f_{\text{esc}}$ .

### 3 SIMULATION SET-UP

Having outlined our fiducial astrophysical model and how the HMF impacts the summary statistics of the 21-cm signal, we now focus on discussing our inference pipeline set-up.

#### 3.1 Parameter inference with SWYFT

When it comes to parameter inference, our goal is to determine the probability distribution of obtaining our model parameters,  $\theta$ , given an observation,  $\mathbf{x}$ . This probability distribution, called the posterior,  $p(\theta | \mathbf{x})$ , is given by Bayes' theorem,

$$p(\theta | \mathbf{x}) = \frac{p(\mathbf{x} | \theta)}{p(\mathbf{x})} p(\theta) \quad (17)$$

where  $p(\mathbf{x} | \theta)$  denotes the likelihood to obtain the observed measurement  $\mathbf{x}$  given the model parameters  $\theta$ ,  $p(\theta)$  constitutes our prior knowledge of the model parameters and  $p(\mathbf{x})$  is the evidence of the data.

Obtaining the posterior can be computationally expensive. Traditionally, to tackle this problem we first assume a functional form for the likelihood and then apply either basic MCMC (e.g. Metropolis et al. 1953; Hastings 1970; Green 1995) or more complex nested

sampling (e.g. Skilling 2004; Skilling 2006; Feroz, Hobson & Bridges 2009; Handley, Hobson & Lasenby 2015; Buchner 2016, 2019), where the latter has the additional benefit of providing an estimate of the evidence which is useful for performing model selection. Fundamentally, these approaches are limited by the inherent assumptions that enter into constructing the functional form of the likelihood and/or how computationally intensive the likelihood calculation becomes.

SBI on the other hand replaces the explicit evaluation of the likelihood with a stochastic simulator that provides the implicit mapping from a set of model parameters  $\theta_i$  to the data  $\mathbf{x}_i$ . These sample-parameter pairs  $[(\mathbf{x}_1, \theta_1), \dots, (\mathbf{x}_N, \theta_N)]$  generated by this stochastic simulator are then used for training the neural network that either estimates the posterior directly, the likelihood or the likelihood-to-evidence ratio (see e.g. Lueckmann et al. 2021, for detailed discussions on benchmarking these different approaches). Our only requirement of this simulator is that it contains a representation of the signal (realization) combined with the noise properties, i.e. modelling or observational uncertainties. The advantage of these approaches is that since they explicitly use the output of the forward-modelled simulations rather than a likelihood evaluation we can accurately characterize the inherent variances (e.g. cosmic variance and instrumental noise) within our simulations and/or our summary statistics.

For this work, we use SWYFT (Miller et al. 2022) which utilizes marginal neural ratio estimation (MNRE; e.g. Durkan, Murray & Papamakarios 2020; Hermans et al. 2021) in order to learn the marginal likelihood-to-evidence ratio rather than the full likelihood-to-evidence ratio of the entire astrophysical parameter set. We denote this marginal likelihood-to-evidence ratio,  $r(\mathbf{x}, \tilde{\theta})$  where  $\tilde{\theta}$  signifies that we are referring to an individual parameter pair (e.g.  $(\theta_i, \theta_j)$ ) rather than the full parameter set,  $\theta$ . We then write this quantity,  $r(\mathbf{x}, \tilde{\theta})$  as

$$r(\mathbf{x}, \tilde{\theta}) \equiv \frac{p(\mathbf{x} | \tilde{\theta})}{p(\mathbf{x})} = \frac{p(\tilde{\theta} | \mathbf{x})}{p(\tilde{\theta})} = \frac{p(\mathbf{x}, \tilde{\theta})}{p(\mathbf{x})p(\tilde{\theta})}. \quad (18)$$

This quantity is equal to the ratio of probability densities for jointly drawn sample-parameter pairs,  $\mathbf{x}, \tilde{\theta} \sim p(\mathbf{x}, \tilde{\theta})$  and marginally drawn pairs  $\mathbf{x}, \tilde{\theta} \sim p(\mathbf{x})p(\tilde{\theta})$ . To determine this ratio, a binary classifier is trained,  $d_{\phi}(\mathbf{x}, \tilde{\theta})$ , where  $\phi$  denotes the network parameters. This discriminates between two hypotheses, whether the sample-parameter pairs are jointly ( $C = 1$ ) or marginally ( $C = 0$ ) drawn with both classes sampled with equal probability. This network is



trained using a binary-cross entropy loss function:

$$L[d_\phi(\mathbf{x}, \tilde{\theta})] = - \int d\mathbf{x} d\tilde{\theta} \{ p(\mathbf{x}, \tilde{\theta}) \log d_\phi(\mathbf{x}, \tilde{\theta}) + p(\mathbf{x}) p(\tilde{\theta}) \log [1 - d_\phi(\mathbf{x}, \tilde{\theta})] \}, \quad (19)$$

which is minimized when  $d_\phi(\mathbf{x}, \tilde{\theta})$  corresponds to the probability of the class labelled  $C = 1$  (i.e. the jointly sampled probability):

$$d_\phi(\mathbf{x}, \tilde{\theta}) = p(C = 1 | \mathbf{x}, \tilde{\theta}) = \frac{p(\mathbf{x}, \tilde{\theta})}{p(\mathbf{x}, \tilde{\theta}) + p(\mathbf{x}) p(\tilde{\theta})} \equiv \sigma[\log r(\mathbf{x}, \tilde{\theta})], \quad (20)$$

which allows the likelihood-to-evidence ratio,  $r$ , to be expressed in terms of the binary classifier,  $d_\phi$ , using the sigmoid function,  $\sigma(y) = [1 + e^{-y}]^{-1}$ .

The key advantage of this approach is that by learning the marginal likelihood-to-evidence ratio, when training this ratio we only need to provide the explicit values for the parameter combinations we are interested in. That is, for parameter inference for an  $M$  dimensional model, we are only required to train  $M$  one-dimensional (1D) and  $M(M-1)/2$  two-dimensional (2D) networks to describe the full marginal posterior distribution. This is because the marginalization over the remaining (nuisance) model parameters is implicitly included within the training data supplied by the stochastic simulator. Thus the variance due to these parameters will be included in the data but is not necessary for training the binary classifier.

### 3.2 Simulated data

In order to perform SBI, we are required to have a stochastic simulator of the cosmic 21-cm signal which includes and characterizes our understanding of the modelling and observational uncertainties. Below, we outline our steps to generate realistic 21-cm data.

Firstly, we use 21CMFAST to generate our 3D realizations of the cosmic 21-cm signal. These are simulated using a comoving box size of  $250^3$  Mpc<sup>3</sup> sampled on  $150^3$  grids using a density field which has been downsampled from a higher resolution box ( $450^3$ ). We construct our 3D 21-cm light-cones spanning from  $z = 25$  down to  $z = 5.2$ . In total, we generate 120 000 realizations of the cosmic 21-cm signal for each of the five different HMFs outlined earlier. Note however, to minimize the total computational overheads we only generate one set of initial conditions per set of 5 different HMFs. However, the astrophysical parameters for each of these HMFs are varied to ensure there is no overlap in the same region of parameter space with the same initial conditions. When constructing the training set it is important to vary the initial conditions as otherwise one would underestimate the sample variance error, which leads to the production of over-confident and biased posterior predictions as discussed in detail in Prelogović & Mesinger (2023, see also Zhao et al. 2022b). Importantly, when constructing our mock observations we use unique random seeds that are mutually distinct from those used within the construction of the training data.

#### 3.2.1 Instrumental noise

In this work, our exploration of the impact of the HMF on our astrophysical parameters is based on a mock observation of the 21-cm PS using the SKA. To embed the instrumental noise characteristics into our mock 21-cm PS, we first split our 3D light-cone along the line of sight into equal comoving distance (250 Mpc) chunks. Next, since radio interferometers only observe the spatial fluctuations in

the signal (zero mean signal), we measure the mean of our individual chunks and remove it from the data cubes.

To mimic the instrumental effects of the SKA on our 3D 21-cm data, we use a modified version of the publicly available PYTHON module 21CMSENSE<sup>7</sup> (Pober et al. 2013, 2014). First, it generates gridded  $uv$ -visibilities based on any instrumental setup, for which we use the antenna configuration outlined in the SKA System Baseline Design document<sup>8</sup> which corresponds to 512 37.5m antennae stations distributed within a 500m core radius. These stations are modelled assuming a system temperature,  $T_{\text{sys}} = 1.1 T_{\text{sky}} + 40$  K with a sky temperature of  $T_{\text{sky}} = 60 \left(\frac{\nu}{300 \text{ MHz}}\right)^{-2.55}$  K (Thompson, Moran & Swenson 2007). We assume a total observing time of 1000 h based on a single six-hour phase-tracked scan of the sky per night. Using these gridded  $uv$ -visibilities, 21CMSENSE then estimates the total thermal noise PS,  $P_N(k)$ ;

$$P_N(k) \approx X^2 Y \frac{\Omega'}{2t} T_{\text{sys}}^2, \quad (21)$$

where  $X^2 Y$  is a conversion between observing bandwidth, frequency and co-moving distance,  $\Omega'$  is a beam-dependent factor derived by Parsons et al. (2014), and  $t$  is the total observing time.

For the purposes of this work, we are not interested in the total thermal noise power, instead we want to directly corrupt the 3D 21-cm data cube. Therefore, following Greig, Ting & Kaurov (2022), we perform the following steps:

- (i) We first 3D Fourier transform the input (simulated) mean removed 21-cm data cube.
- (ii) We then filter this cube using the gridded  $uv$ -visibilities for the SKA computed by 21CMSENSE. Cells with finite  $uv$ -coverage are multiplied by unity, all others are set to zero.
- (iii) At each cell we then determine the amplitude of the thermal noise,  $P_N(k_x, k_y, k_z)$ , using equation (21) where  $k_x$  and  $k_y$  correspond to the two transverse (on sky) directions and  $k_z$  is the line-of-sight direction.
- (iv) We then add random noise (zero mean with variance based on the power spectrum amplitude in the cell) to each cell to mimic the effect of thermal noise.
- (v) Finally, we then 3D inverse Fourier transform back to obtain our noisy 21-cm data.

#### 3.2.2 Foreground avoidance

Unfortunately, the  $uv$  visibilities of radio interferometers are frequency dependent meaning that line-of-sight (frequency dependent) power can bleed into the transverse (frequency independent) Fourier modes. This ‘feature’ manifests as a relatively well-defined contaminated ‘wedge’ in cylindrical 2D Fourier space (Datta, Bowman & Carilli 2010; Morales et al. 2012; Parsons et al. 2012; Trott, Wayth & Tingay 2012; Vedantham, Udaya Shankar & Subrahmanyan 2012; Thyagarajan et al. 2013; Liu, Parsons & Trott 2014a,b; Thyagarajan et al. 2015a,b; Pober et al. 2016; Murray & Trott 2018). Although it is theoretically possible to mitigate or ‘clean’ these contaminated modes (see e.g. Chapman & Jelić 2019 for a review, or by using machine learning Gagnon-Hartman et al. 2021) in this work we take the conservative approach of cutting all contaminated ‘wedge’

<sup>7</sup><https://github.com/jpober/21cmSense>

<sup>8</sup>[http://astronomers.skatelescope.org/wp-content/uploads/2016/09/SKA-TEL-SKO-0000422.02\\_SKA1\\_LowConfigurationCoordinates-1.pdf](http://astronomers.skatelescope.org/wp-content/uploads/2016/09/SKA-TEL-SKO-0000422.02_SKA1_LowConfigurationCoordinates-1.pdf)

modes from our 21-cm light-cones. That is, we adopt a foreground-avoidance observing strategy, whereby we only measure our 21-cm PS using ‘clean’ Fourier modes above this ‘wedge’.

The boundary confining this foreground ‘wedge’ in 2D Fourier space is given by

$$k_{\parallel} = mk_{\perp} + b, \quad (22)$$

where  $k_{\parallel}$  and  $k_{\perp}$  are the line of sight and transverse Fourier modes,  $b$  is an additive buffer which we assume to be  $\Delta k_{\parallel} = 0.1 h \text{ Mpc}^{-1}$  which accounts for bleeding of noise extending beyond the horizon limit and  $m$  is the gradient of this boundary given by

$$m = \frac{D_C H_0 E(z) \sin(\theta)}{c(1+z)}. \quad (23)$$

This boundary depends on the comoving distance,  $D_C$ , the Hubble constant,  $H_0$ , cosmological factor  $E(z) = \sqrt{\Omega_m(1+z)^3 + \Omega_\Lambda}$  and  $\sin(\theta)$  denotes the viewing angle of the telescope, for which we conservatively take as  $\theta = \pi/2$  (i.e. a zenith pointing observation).

Including the imprint of foreground avoidance in our inference pipeline relies on an additional step to those outlined above. After 3D Fourier transforming our input 3D 21-cm data cube, we first zero all modes that fall below this foreground ‘wedge’ before adding the thermal noise for all modes above the wedge.

## 4 RESULTS WHEN INCORRECTLY ASSUMING A SPECIFIC HMF

Our first investigation is to quantify the bias in our inferred astrophysical parameters based on the assumption of an incorrectly chosen HMF model. Using our mock 21-cm observation assuming a ST HMF, we split the 3D 21-cm light-cone into ten 21-cm PS spanning from  $z = 5.7$  to  $z = 18.1$  (see e.g. Fig. 3). Although the SKA is expected to observe down to 50 MHz ( $z \sim 27.8$ ), at these redshifts the thermal noise significantly dominates over the 21-cm signal. Therefore, for computational ease we restrict our data to a lower redshift ( $z \leq 18$ ). For each 21-cm PS, we only consider Fourier modes between  $k = 0.1 \text{ Mpc}^{-1}$  and  $k = 1.0 \text{ Mpc}^{-1}$ , which leads to six Fourier modes per 21-cm PS. This lower limit loosely corresponds to when cosmic variance begins to dominate over the 21-cm signal whereas the limit at  $k = 1.0 \text{ Mpc}^{-1}$  is when shot-noise in the numerical modelling becomes dominant (Greig & Mesinger 2015).

Following Saxena et al. (2023), to obtain our marginal posterior distributions for our astrophysical parameters we express our 21-cm PS data as a single, 1D array (i.e. 60 data points). This data is then the input layer to a three-layered fully connected neural network consisting of 256 neurons. The network is trained in batches of size 64, with an initial learning rate of  $10^{-3}$  that is decayed by 0.95 after each epoch. For a single HMF, the training of all the MNRE networks for our astrophysical model takes  $\sim 1$  hour using a single Nvidia A100. Once we have our MNRE networks, parameter inference only takes a few minutes. One of the significant advantages of MNRE is that we can more readily assess the performance of our trained network using the network convergence (Cole et al. 2022). We provide an example of this in Appendix A.

### 4.1 Mock observation using only the 21-cm PS

In Fig. 5, we present our marginalized posteriors for our 1000 h mock observation of the 21-cm PS with the SKA, which assumes our fiducial parameter set and a ST HMF. Along the diagonal, we provide the 1D marginalized posteriors whereas above and below the diagonal

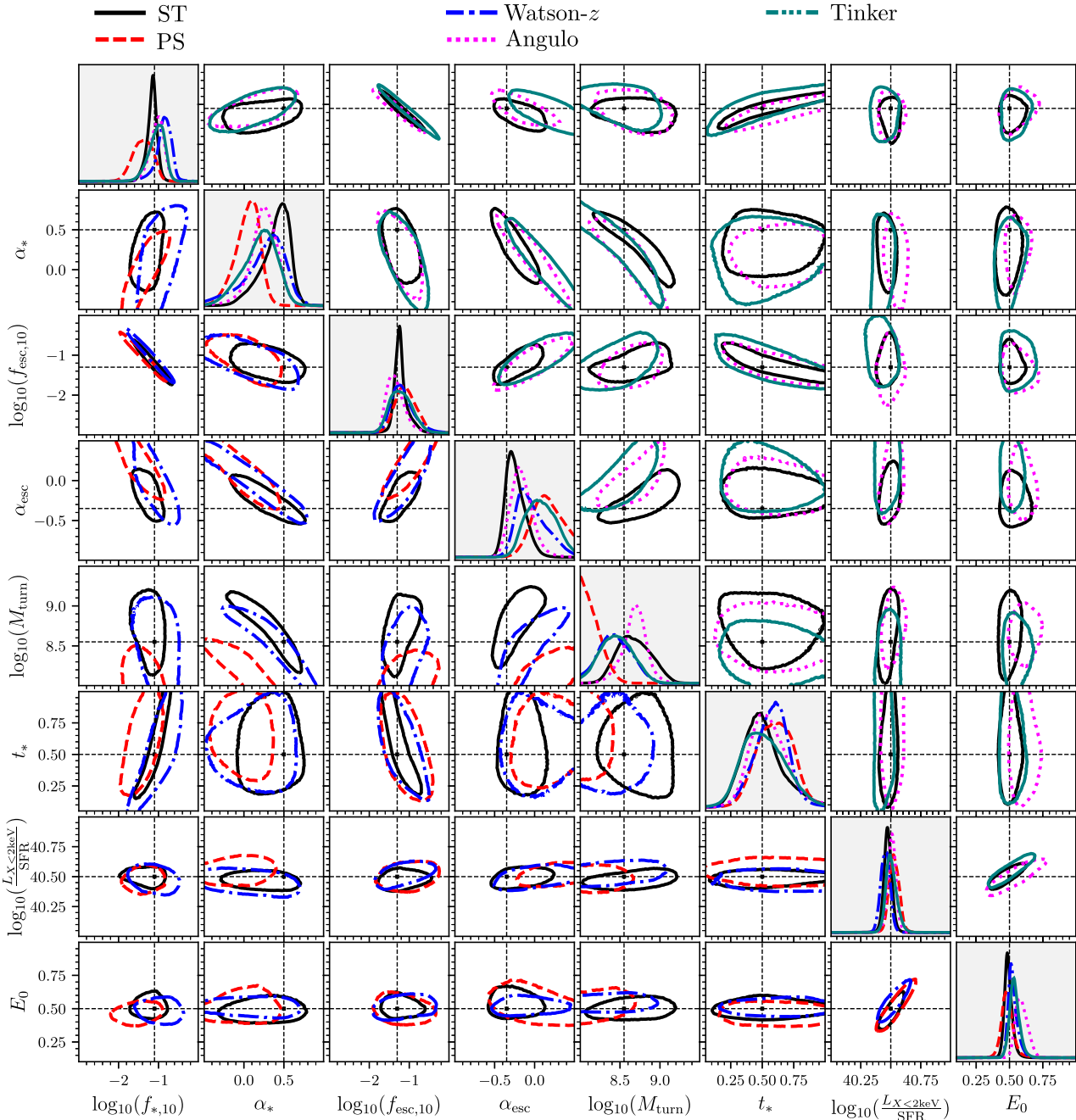
we provide the 95th percentile joint marginalized posteriors. Using our mock observation, we train SWYFT on our forward-modelled simulation data fixing the HMF for the same mock.<sup>9</sup> The coloured curves correspond to assuming ST (solid, black), PS (red, dashed), Watson- $z$  (blue, dot-dashed), Angulo (magenta, dotted), and Tinker (teal, triple-dot-dashed). Finally, in Table 2 we summarize the 68th marginalized percentiles for each model.

Focusing first on the ST HMF results, which serves as our baseline for the relative amplitudes of the marginalized uncertainties, we recover tight (and unbiased) constraints on our model parameters. Further, the relative uncertainties are broadly consistent with the direct MCMC approaches using 21CMMC by Park et al. (2019) and Greig, Mesinger & Koopmans (2020), albeit for a slightly different model. Therefore, as was observed by Saxena et al. (2023), who performed a direct comparison between the approaches, SWYFT performs extremely well at obtaining robust posteriors using the 21-cm signal.

Immediately evident from Fig. 5 are several strong biases in the inferred astrophysical parameters owing to the incorrect HMF being assumed for our mock observation. These biases are strongest for the UV galaxies parameters, with the X-ray parameters being relatively unaffected. However, this is not too surprising given the role of the star-forming galaxies throughout the entirety of reionization. The SFR both controls the number of UV photons produced for ionizing the IGM while also setting the amplitude of the production rate of X-rays which subsequently heat the IGM during the epoch of heating. Although the X-ray emissivity is proportional to the product of the X-ray luminosity and SFRD (see equation 10), the dominant role of the latter during reionization ensures that the X-ray parameters are unaffected by the choice of HMF despite the increasingly larger differences in the HMFs for increasing redshift. This same behaviour was also observed within the global 21-cm signal analysis by Mirocha et al. (2021). Further contributing to the reduced impact on the X-ray parameters will be the relatively larger observational uncertainties which increase for increasing redshift, where the 21-cm signal becomes more sensitive to X-ray parameters. Below, we perform an in-depth analysis of the various trends in our inferred parameters to gain further insight into the role of the HMF.

The largest biases occur for the PS model relative to our mock ST observation. For example, the inferred value of  $M_{\text{turn}}$  peaks at the edge of our prior region.<sup>10</sup>  $M_{\text{turn}} = 10^8 M_{\odot}$ , implying that to match the mock observation the PS HMF must produce many more low-mass galaxies. This requirement is primarily set by the X-ray heating epoch, where from Fig. 2, the absorption trough is delayed by  $\Delta z \sim 2$  relative to the ST model. Therefore, to shift to a comparable timing in order to have the 21-cm PS achieve similar amplitudes at the correct redshifts, the PS model requires many more haloes at earlier times due to the fact that the amplitude of the PS HMF drops away significantly to increasing redshift (see Fig. 1). However, the PS HMF also rapidly produces star-forming galaxies during the EoR, as seen by the sharper gradient in the reionization history. Therefore, simply

<sup>9</sup>Note, when initially performing our parameter inference we also trialed the SBI approach developed by Prelogović & Mesinger (2023), namely the conditional masked-autoregressive flow model to learn the actual likelihood before utilizing it within an MCMC framework for inference. For completeness in Appendix B we verified the consistency of these two approaches, however, for the remainder of this work we only use SWYFT for computational efficiency.  
<sup>10</sup>Note, this lower bound is set by only considering atomically cooled galaxies. In order to reduce this boundary, we would need to also consider a secondary population of molecularly cooled galaxies (e.g. Qin et al. 2020, 2021a).



**Figure 5.** The recovered 1D and 2D marginalized posteriors on our astrophysical parameters, assuming a mock 1000 h observation of the 21-cm PS assuming foreground avoidance with the SKA (see Section 3.2 for further details). For this our mock observation assumes a ST HMF (Sheth et al. 2001), while our forward modelled simulations assume a different fixed HMF. In particular, we consider PS (Press & Schechter 1974; red-dashed), Tinker (Tinker et al. 2010; teal triple dot-dashed), Angulo (Angulo et al. 2012; magenta-dotted) and the redshift dependant Watson (Watson et al. 2013; blue-dot-dashed) HMFs. The diagonal panels correspond to the marginalized 1D posteriors whereas the off-diagonal contours correspond to the 95th percentile joint marginalized posteriors. The vertical and horizontal black dashed lines denote our fiducial astrophysical parameter set. Note, for ease of visualization, we split the 2D posteriors above (Angulo and Tinker) and below (PS and Watson-z) the diagonal relative to the fiducial ST HMF.

dropping  $M_{\text{turn}} = 10^8 M_{\odot}$  would cause the EoR to occur much earlier than for the ST model. To prevent this, we are required to delay (or suppress) the escape of the UV photons from these star-forming galaxies. However, since we require star formation to produce X-rays, the dominant way we can suppress ionizations during the EoR is through modifications to  $f_{\text{esc}}$  (i.e. prevent them escaping the host galaxies). For our mock observation, we have  $\alpha_{\text{esc}} = -0.35$  implying

higher escape fractions for lower mass galaxies. Instead, if we assume a PS HMF, we recover  $\alpha_{\text{esc}} = 0.12$ , which suppresses the escape of ionizing photons by a factor of 10. Note, we also see a reduction in amplitude in  $\log_{10}(f_{*,10})$  of  $\sim 20$  per cent, which corresponds to an absolute reduction in star formation of  $\sim 70$  per cent. In summary, the rapidly evolving amplitude and shape of the PS HMF drives these relatively extreme biases for the astrophysical parameters.

**Table 2.** A summary of the recovered astrophysical parameter constraints plus 68th percentiles when we vary the underlying HMF in our inference pipeline. For this, we consider a mock observation of the 21-cm PS, assuming foreground wedge avoidance and a 1000 h observation with the SKA. Note, for our mock observation we assume a ST HMF (Sheth et al. 2001). Additionally, we consider the recovered posteriors for our astrophysical parameters once we fold in the additional constraining power from UV LFs (see the text for further details).

	$\log_{10}(f_{*,10})$	$\alpha_*$	$\log_{10}(f_{\text{esc},10})$	$\alpha_{\text{esc}}$	$\log_{10}(M_{\text{turn}})$ ( $M_{\odot}$ )	$t_*$	$\log_{10}\left(\frac{L_{X<2\text{keV}}}{\text{SFR}}\right)$ ( $\text{erg s}^{-1} M_{\odot}^{-1} \text{yr}^{-1}$ )	$E_0$ (keV)
Mock Observation	-1.1	0.5	-1.30	-0.35	8.55	0.5	40.5	0.5
21-cm PS only								
ST	$-1.14^{+0.10}_{-0.13}$	$0.48^{+0.13}_{-0.19}$	$-1.25^{+0.12}_{-0.14}$	$-0.29^{+0.14}_{-0.08}$	$8.60^{+0.27}_{-0.15}$	$0.48^{+0.14}_{-0.15}$	$40.47^{+0.03}_{-0.03}$	$0.49^{+0.03}_{-0.02}$
PS	$-1.37^{+0.23}_{-0.26}$	$0.11^{+0.11}_{-0.18}$	$-1.13^{+0.27}_{-0.24}$	$0.12^{+0.22}_{-0.15}$	$8.00^{+0.20}_{-0.00}$	$0.59^{+0.16}_{-0.13}$	$40.53^{+0.04}_{-0.04}$	$0.48^{+0.04}_{-0.04}$
Tinker	$-0.98^{+0.16}_{-0.22}$	$0.25^{+0.19}_{-0.22}$	$-1.28^{+0.31}_{-0.23}$	$0.03^{+0.22}_{-0.16}$	$8.44^{+0.21}_{-0.20}$	$0.46^{+0.20}_{-0.13}$	$40.49^{+0.05}_{-0.03}$	$0.53^{+0.04}_{-0.03}$
Angulo	$-0.98^{+0.16}_{-0.21}$	$0.25^{+0.17}_{-0.14}$	$-1.42^{+0.23}_{-0.18}$	$-0.19^{+0.15}_{-0.12}$	$8.70^{+0.14}_{-0.14}$	$0.51^{+0.16}_{-0.12}$	$40.51^{+0.03}_{-0.03}$	$0.58^{+0.05}_{-0.05}$
Watson-z	$-0.85^{+0.18}_{-0.15}$	$0.37^{+0.19}_{-0.27}$	$-1.25^{+0.27}_{-0.24}$	$-0.15^{+0.24}_{-0.12}$	$8.38^{+0.25}_{-0.16}$	$0.61^{+0.10}_{-0.15}$	$40.46^{+0.03}_{-0.04}$	$0.51^{+0.03}_{-0.03}$
21-cm PS + UV LFs								
ST	$-1.11^{+0.13}_{-0.13}$	$0.45^{+0.08}_{-0.09}$	$-1.20^{+0.13}_{-0.11}$	$-0.30^{+0.05}_{-0.05}$	$8.47^{+0.16}_{-0.12}$	$0.58^{+0.16}_{-0.12}$	$40.49^{+0.03}_{-0.02}$	$0.49^{+0.04}_{-0.03}$
PS	$-0.95^{+0.14}_{-0.21}$	$0.52^{+0.14}_{-0.08}$	$-1.40^{+0.15}_{-0.19}$	$-0.17^{+0.04}_{-0.06}$	$8.00^{+0.07}_{-0.00}$	$0.59^{+0.20}_{-0.25}$	$40.50^{+0.04}_{-0.03}$	$0.57^{+0.08}_{-0.05}$
Tinker	$-0.92^{+0.13}_{-0.17}$	$0.48^{+0.08}_{-0.08}$	$-1.41^{+0.19}_{-0.14}$	$-0.23^{+0.04}_{-0.05}$	$8.31^{+0.11}_{-0.15}$	$0.43^{+0.23}_{-0.14}$	$40.52^{+0.03}_{-0.03}$	$0.52^{+0.05}_{-0.05}$
Angulo	$-1.15^{+0.11}_{-0.13}$	$0.45^{+0.08}_{-0.08}$	$-1.32^{+0.18}_{-0.09}$	$-0.20^{+0.05}_{-0.07}$	$8.27^{+0.16}_{-0.15}$	$0.65^{+0.16}_{-0.17}$	$40.50^{+0.04}_{-0.03}$	$0.52^{+0.04}_{-0.04}$
Watson-z	$-1.03^{+0.12}_{-0.13}$	$0.34^{+0.08}_{-0.07}$	$-1.33^{+0.16}_{-0.16}$	$-0.15^{+0.06}_{-0.04}$	$8.37^{+0.18}_{-0.09}$	$0.64^{+0.11}_{-0.20}$	$40.47^{+0.04}_{-0.03}$	$0.49^{+0.05}_{-0.04}$

Generally speaking for the remaining HMFs, the relative biases are less extreme, owing to their similar and more modest evolutions in the overall HMF amplitude as well as the power-law slope and exponential decay for increasing halo mass. As one would anticipate from Fig. 2, both Tinker and Watson-z recover similar behaviour. For these we have a delay in EoR heating of  $\Delta z \sim 1$ , thus we recover a preference for a lower  $M_{\text{turn}}$  by  $\sim 2$  per cent (logarithmic), which albeit seemingly small corresponds to a  $\sim 1\sigma$  reduction or a  $\sim 40$  per cent lower halo mass. Additionally, we prefer a  $\sim 15 - 30$  per cent increase in  $\log_{10}(f_{*,10})$  corresponding to an absolute increase in the star formation of  $\sim 50 - 100$  per cent. Note, although these differences appear quite large for relatively similar models, there is a strong degeneracy between  $\log_{10}(f_{*,10})$  and  $\alpha_{\text{esc}}$ , thus the higher  $\log_{10}(f_{*,10})$  for the Watson-z model is compensated for by a correspondingly lower  $\alpha_{\text{esc}}$ . These same trends also follow for the Angulo HMF, albeit in the opposite direction. For an Angulo HMF, reionization occurs slightly earlier, thus we require a shift to a slightly higher  $\log_{10}(M_{\text{turn}}) \sim 8.70$  to slightly delay it (fewer low-mass galaxies). However, this has the consequence of marginally delaying the onset of X-ray heating, which can be compensated for by increasing  $\log_{10}(f_{*,10})$ , but equally kept in balance by a slightly higher  $E_0$  (ensuring X-ray heating is not too efficient, albeit with a larger uncertainty). To ensure the increasing  $\log_{10}(f_{*,10})$  also does not result in early reionization a higher  $\log_{10}(f_{\text{esc},10})$  is preferred.

In summary, the relative biases due to the choice of assumed HMF is governed by ensuring the resultant 21-cm PS matches the mock observation both during the EoR and during the X-ray heating epoch. Overall, depending on the choice of HMF, we can recover differences between the input and recovered constraints in  $\log_{10}(f_{*,10})$  of  $\pm 30$  per cent, up to  $\sim 20$  per cent for  $\alpha_*$ ,  $\pm 15$  per cent for  $\log_{10}(f_{\text{esc},10})$ , up to  $\sim 7$  per cent for  $\log_{10}(M_{\text{turn}})$  and up to  $\sim 20$  per cent for  $t_*$ . In terms of the relative uncertainties from our mock ST observation, these correspond to up to 2, 2, 1, 3,  $0.5\sigma$ , respectively. Note, throughout this work we define these relative biases as the separation between the maximum *a*-posteriori values for each individual parameter to the fiducial values of the ST mock observation divided by the recovered  $1\sigma$  values when assuming the correct ST HMF. By far the largest impact is on  $\alpha_{\text{esc}}$  which controls

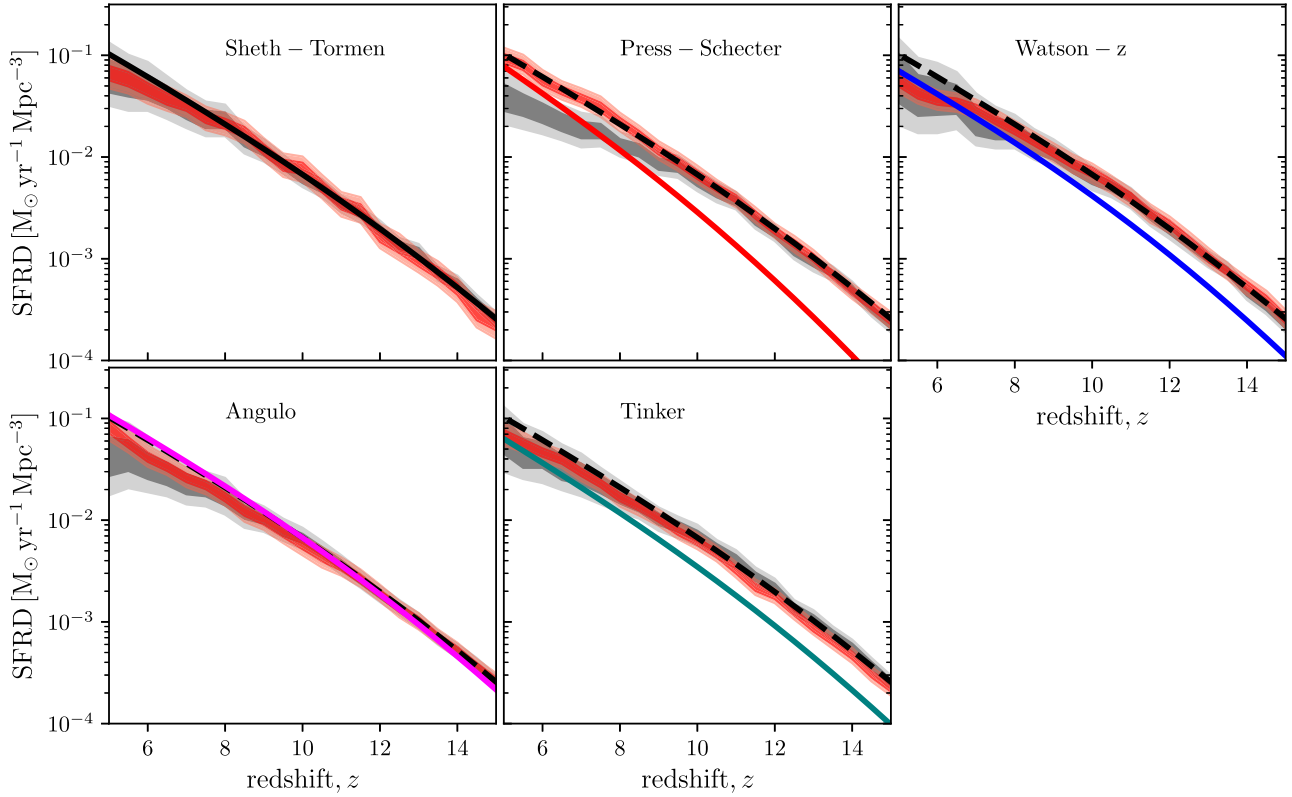
the mass dependence of the escape fraction. Often we require more X-rays earlier, which can be achieved through an increased SFR, with the consequence of causing reionization to occur too early. Therefore the only way we can limit ionizations during the EoR is through,  $f_{\text{esc}}$  and principally through the mass dependence,  $\alpha_{\text{esc}}$ . For  $\alpha_{\text{esc}}$  we can recover biases up to  $\sim 150$  per cent, which roughly corresponds to  $\sim 3 - 4\sigma$ .

As noted earlier, the evolution of the 21-cm signal and summary statistics for the Angulo HMF model are extremely similar to our mock ST HMF model (see Figs 2 and 3). Therefore, we can use the differences between these two models as a guide to the impact of even relatively small differences between the assumed HMF. Of our eight astrophysical parameters, four are constrained to values either right at the edge or beyond the 68th percentile boundaries. Most notably,  $\alpha_*$  is  $\sim 2$  times lower than the input value. Therefore, even relatively modest differences between the underlying HMF can cause biases in the inferred astrophysical parameters at or beyond the inferred 68th percentile limits.

## 4.2 Star formation rate density

In the previous section, we established that the choice of HMF only impacts the UV properties of the galaxies, with the X-ray parameters left relatively unaffected (although we do observe a fairly minor bias in  $E_0$  for the Angulo model). The relative biases in these UV galaxy parameters are driven by each model attempting to match the 21-cm PS evolution over cosmic history to that of the mock ST HMF observation. The most relevant quantity tied to the evolution of the 21-cm signal with redshift is the SFRD. First, this dictates when the 21-cm signal decouples from the background CMB during the Ly  $\alpha$  heating epoch. It then controls the production rate of X-ray photons (see equation 10) before finally driving reionization (in combination with the escape fraction,  $f_{\text{esc}}$ ; see equation 14).

To more readily demonstrate this, in Fig. 6 we provide the recovered posteriors on the mean SFRD for each assumed HMF given our mock 21-cm PS observation with the SKA. We provide the recovered marginalized 68th (95th) percentiles as dark (light) shaded regions for each HMF along with the mean SFRD from



**Figure 6.** The marginalized posteriors on the mean SFRD following our SBI pipeline for our mock 21-cm PS observation with the SKA. Each panel highlights the 68th (dark) and 95th (light) marginalized posteriors for an assumed HMF. The grey contours correspond to the marginalized posteriors using only the 21-cm PS and the red contours correspond to the posteriors when using both the 21-cm PS and the UV LFs. The black dashed curve in each panel is the mean SFRD for the mock observation (assuming a Sheth–Tormen HMF). The coloured curves in each panel correspond to the expected SFRD assuming the same astrophysical parameter set as our mock observation but with a different HMF.

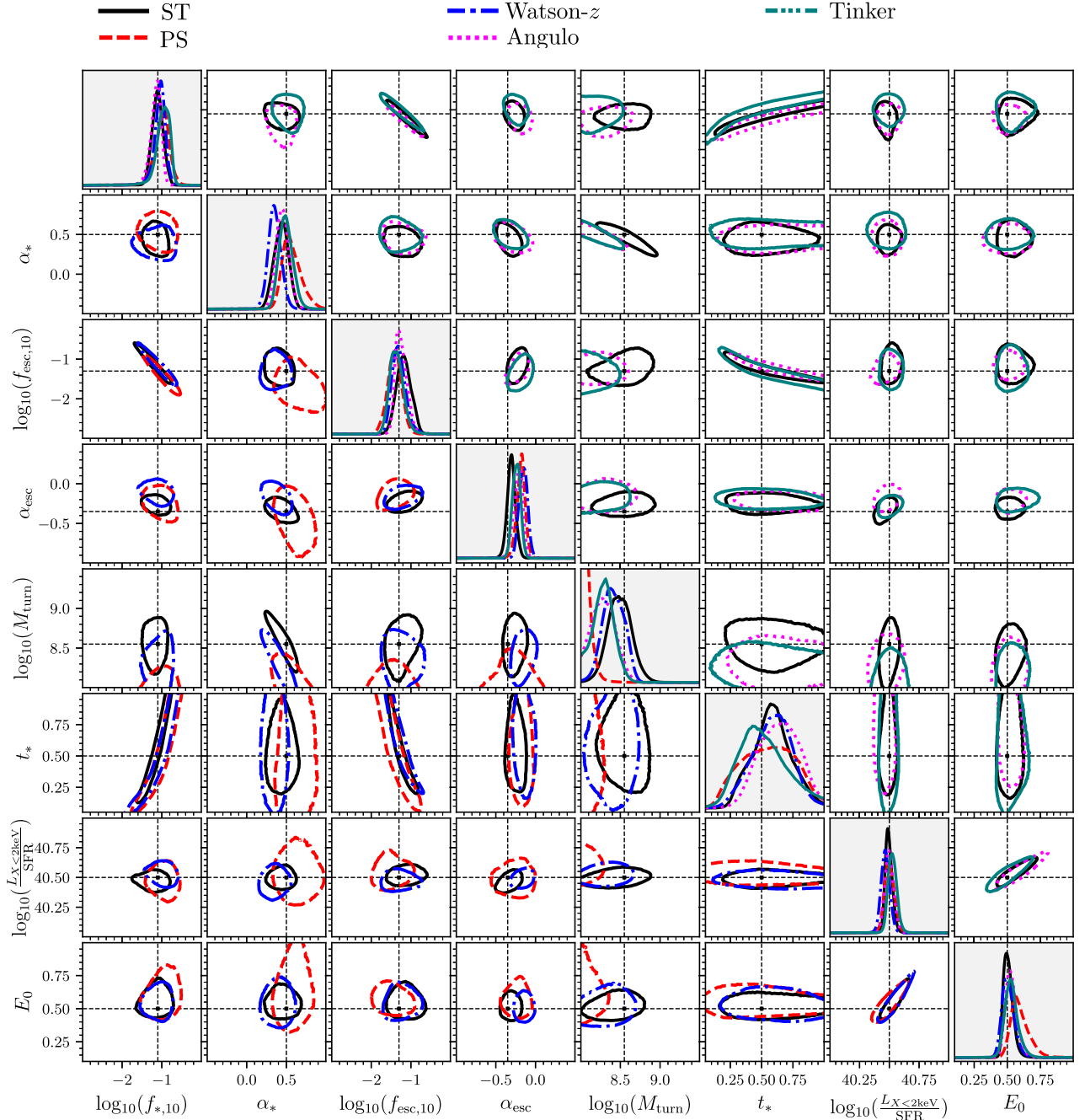
the mock observation (black dashed curve). For reference, we also provide the mean SFRD assuming the same astrophysical parameters as our mock observation, but with each assumed HMF to visualize the necessary shift in the SFRD. Within SWYFT this is simply achieved by modifying the end-points of the fully connected network architecture used for estimating the marginalized likelihood-to-evidence ratio. Retaining the same 21-cm PS data as input, we simply replace our astrophysical parameters with the SFRD values as a function of redshift. This is allowed as our training set inherently varies the underlying astrophysical parameters. This retraining of the network is only necessary as we are using MNRE, where we recover the marginalized likelihood-to-evidence ratio. If instead we obtained the full likelihood-to-evidence ratio we could simply sample the full likelihood-to-evidence ratio within an MCMC framework to obtain the posteriors on the SFRD.

As expected, by assuming an incorrect HMF, the inferred astrophysical parameters that we recovered in the previous section are biased to ensure we reproduce the correct SFRD from our mock 21-cm observation. This is readily demonstrated by the shift away from the expected SFRD for each HMF (coloured curves) to the tight posteriors around the SFRD of the mock (black dashed curves). Adopting a PS, Watson- $z$  or Tinker HMF under predicts the mean SFRD, therefore the inferred astrophysical parameters we recovered previously were biased to increase the overall SFRD. For example, we inferred lower  $M_{\text{turn}}$  for these models and in general an increased  $f_*$  which is consistent with this trend. For the

Angulo HMF, it marginally over predicts and thus we inferred biased astrophysical parameters which reduce the overall stellar output, for example our observed increase in  $M_{\text{turn}}$  (i.e. less star-forming galaxies).

For decreasing redshift,  $z < 8$ , the inferred mean SFRD tends to drop below the expectation from the mock observation (black solid and dashed curves). However, this is simply due to the strong degeneracy between star formation,  $f_*$ , and the escape fraction,  $f_{\text{esc}}$  (see e.g. Fig. 5). Here, when considering the SFRD we are only considering  $f_*$ , whereas reionization is driven by the product of  $f_*$  and  $f_{\text{esc}}$  (equation 14). For example, even when we assumed the correct ST HMF, we observe a slight decrease in the inferred mean SFRD relative to the mock observation. From Table 2, we recovered a slightly reduced  $\log_{10}(f_{*,10})$  ( $-1.14$  compared to the input  $-1.1$ ) and correspondingly higher  $\log_{10}(f_{\text{esc},10})$  ( $-1.25$  compared to the input  $-1.30$ ) along with a less negative slope for  $\alpha_{\text{esc}}$  ( $-0.29$  compared to the input  $-0.35$ ), which results in the slight decrease in the inferred SFRD during the EoR. This same trend additionally occurs for the other HMFs albeit with a larger amplitude owing to the different properties of the underlying HMF. For the PS HMF, the discrepancy is much larger, which corresponds to the preference for an  $M_{\text{turn}}$  right at the edge of our prior range which translates to much larger biases in the other galaxy parameters.

In addition to this relative downturn in the amplitude of the inferred mean SFRD at  $z < 8$  we also observe an increase in the width of the corresponding posteriors. This increasing width is also due



**Figure 7.** The same as Fig. 5 except also including the UV LFs at  $z = 6, 7, 8$  and  $10$  as observational priors.

to the degeneracy between  $f_*$  and  $f_{\text{esc}}$ . At high redshifts, prior to reionization, the relative amplitude of the 21-cm signal (and statistics such as the PS) are set by X-ray heating, which are not sensitive to  $f_{\text{esc}}$  (their effective escape fraction is instead controlled by  $E_0$ ). As a result, the posteriors at higher redshift, where the 21-cm PS is at its largest amplitude are tighter.

### 4.3 Mock 21-cm PS plus UV LF observation

Thus far we have only explored the relative bias in our inferred parameters for a mock 21-cm PS observation. However, when inferring observations of the 21-cm signal, independent constraining power

is available through complimentary observations. Most notably, through the addition of UV LFs (Park et al. 2019). As outlined in Section 2.6.3, UV LFs serve to break the degeneracy between  $f_*$  and  $f_{\text{esc}}$  by providing independent constraints on  $f_*$ . Therefore, in this section, we explore the impact of our choice of HMF on our inferred parameters when combining our mock 21-cm PS measurement with the SKA and existing UV LF measurement. Namely, we consider UV LFs at  $z = 6, 7, 8$  and  $10$ .

In Fig. 7, we provide the marginalized posteriors for our 1000 h mock observation of the 21-cm PS from the SKA combined with our 4 UV LFs. Along the diagonal, we provide the 1D marginalized posteriors whereas above and below the diagonal we provide the 95th

percentile joint marginalized posteriors. The coloured curves denote the inferred astrophysical parameters for different assumed HMFs fit to the same mock observation (ST HMF). Again, we provide the marginalized 68th percentiles in Table 2.

Importantly, for SBI we require a simulator of the forward-modelled data which includes the dominant sources of uncertainty. In 21CMFAST, the output UV LFs are computed analytically. Therefore, to include noise in our simulated UV LFs we simply add random noise,  $\mathcal{N}(0, \sigma)$  into each corresponding  $M_{UV}$  bin according to the relative observational uncertainty,  $\sigma(M_{UV})$ . Fundamentally, this assumes each  $M_{UV}$  bin is an independent measurement, which should be a reasonable assumption. In Appendix C we verify this approach reproduces the expected marginalized posteriors in comparison to a simple MCMC when considering only UV LFs. Further, our mock astrophysical model is based on the results of Qin et al. (2021b), which deviates from the best-fitting model of Park et al. (2019) which is constrained against the same UV LFs used in this work. Therefore, in order to avoid biasing our results due to the different astrophysical parametrizations, we simply rescale the observed UV LFs by the difference between our fiducial parameter set and that of Park et al. (2019) to ensure our mock will be consistent with the input UV LFs. Since we are focused on quantifying the biases owing to the choice of HMF, and not constraining against real-world data, this approach is justified.

Firstly, considering our fiducial model (ST HMF), one can clearly see the impact that the UV LFs have on improving the constraining power on the UV galaxy properties. Primarily, we see significant improvements in  $f_{*,10}$ ,  $\alpha_*$ , as these are the two most strongly constrained galaxy parameters from the UV LFs (see e.g. Fig. C1). Further, as this is independent information, the inclusion of UV LFs also serve to break the degeneracies between  $f_*$  and  $f_{esc}$  as can be clearly seen both by the notably reduced joint posterior volumes and the orientation (change in slope) of the posteriors contours. Quantitatively, following the inclusion of the UV LFs we observe a reduction in the uncertainties of  $\sim 25$  per cent for  $\log_{10}(f_{esc,10})$  with the most significant improvements in both  $\alpha_{esc}$  and  $\alpha_*$  of  $\sim 50$  per cent. These improvements are consistent with those seen by Park et al. (2019).

Relative to Fig. 5, the maximum *a*-posteriori values of the astrophysical parameters for the different incorrectly chosen HMFs has been reduced relative to the true fiducial parameters. Primarily, this reduction stems from the shrinking posterior volumes that come with the breaking of the astrophysical parameter degeneracies following the inclusion of the UV LFs. This now significantly limits the available posterior space to achieve the correct underlying SFRD to match that of the mock observation. To demonstrate this, we again use SWYFT to obtain our marginalized posteriors on the SFRD when combining UV LFs with our mock 21-cm observation. In Fig. 6, the dark (light) red shaded regions denote the 68th (95th) percentiles for our joint UV LFs and 21-cm mock observation. Quite clearly, one can see significant reductions in the relative widths of the posteriors along with less of a significant downturn in the marginalized posteriors during reionization (e.g. at  $z \lesssim 8$ ). The removal of this downturn, along with the notably reduced posterior widths corresponds to the breaking of the degeneracy between  $f_*$  and  $f_{esc}$ . Now, with  $f_*$  strongly constrained by the UV LFs, there is now less posterior volume with which  $f_{esc}$  can achieve reionization comparably to the mock observation.

For the PS HMF, we again prefer  $M_{turn}$  right at the edge of our prior range ( $10^8$ ); however, this is now inconsistent with our assumed value of 8.55 by  $\sim 8\sigma$ . Previously to compensate for this notably reduced  $M_{turn}$  we required a reduced  $f_{*,10}$  and an  $f_{esc}$  parametrization which

significantly reduced the escape fraction in low-mass galaxies. By including UV LFs, we can no longer significantly reduce the stellar component as we would incorrectly match the observed UV LFs, thus we can only modify  $f_{esc}$ . Because of these limitations, we now prefer a considerably different  $f_{esc}$ , with lower  $\log_{10}(f_{esc,10})$ ,  $-1.4$  compared to  $-1.13$  previously, and a decreasing halo mass dependence. As a consequence of this enforced higher SFR, we would overproduce the number of X-rays; however, we now observe a bias in  $E_0$  shifting to a higher value ( $\sim 1.5\sigma$ ) to reduce the number of X-rays capable of escaping the host galaxy.

Like previously, for the remaining HMFs the behaviour of the biases are less extreme. For the Tinker and Watson- $z$  HMFs, we were required to increase the relative number of UV photons to ensure reionization could occur sooner, resulting in an overall increase of  $f_*$  (primarily for lower mass haloes) and a decrease in  $M_{turn}$ . Following the inclusion of the UV LFs, we are now restricted by the number of UV photons we can produce, which is compensated for by a slight further reduction in  $M_{turn}$  to increase the number of sources (especially for the Tinker HMF). Since we are no longer producing as many stars per stellar baryon for the low-mass galaxy end, we must equivalently increase  $f_{esc}$  for these galaxies to ensure enough of these UV photons are entering the IGM to allow reionization to occur earlier. For the Angulo HMF model, previously we required increasing  $M_{turn}$  and lowering  $f_{esc}$  to slightly delay the onset of reionization and equally a higher  $f_*$  to produce a sufficient number of X-rays. Including UV LFs restricts our ability to increase  $f_*$ , thus instead we are required to reduce  $M_{turn}$ . This change to a lower  $M_{turn}$  results in a notable change in the  $f_*$  mass dependence (requiring fewer stars for lower mass galaxies) whose change is balanced by an increase in  $f_{esc}$  to ensure the timing of reionization is correct.

In summary, the relative amplitudes of the biases in the astrophysical parameters have drastically reduced as we no longer have the freedom to vary  $f_*$  to any value/mass dependence as it is now strongly constrained by the UV LFs. As a consequence, this results in a significant reduction in the available posterior volumes for our model parameters that limits the corresponding biases. Nevertheless, although the relative amplitudes of the biases away from the fiducial parameters have reduced, quantitatively strong biases remain owing to the significantly reduced marginalized uncertainties provided by the UV LFs. For example, we now recover differences in  $\log_{10}(f_{*,10})$  of  $\sim 20$  per cent,  $\sim 30$  for  $\alpha_*$ ,  $\sim 10$  per cent for  $\log_{10}(f_{esc,10})$ , up to  $\sim 7$  per cent for  $\log_{10}(M_{turn})$  and  $\sim 30$  per cent for  $t_*$ . Relative to the marginalized uncertainties these correspond to up to 1.5, 2, 1, 7,  $1\sigma$ , respectively. For  $\alpha_{esc}$ , we recover much smaller biases of  $\Delta\alpha_{esc} \sim 0.20$  compared to  $\Delta\alpha_{esc} \sim 0.50$  previously; however, this still corresponds to  $\sim 3-4\sigma$  bias owing to the factor of  $\sim 2-3$  reduction in the corresponding marginalized uncertainties.

Again, we can use the Angulo model as an illustrative example of the relative biases for even modest differences in the HMF away from the mock observation. Following the inclusion of the UV LFs, now only two parameters are in excess of the 68th marginalized uncertainties relative to the fiducial model compared to the four previously when only considering the 21-cm PS. However, the two inconsistent parameters  $\alpha_{esc}$  and  $M_{turn}$  are more significantly biased (owing to the restriction of  $f_*$  provided by the UV LF data). For example,  $\alpha_{esc}$  is recovered to  $\sim -0.2$  which is now  $\sim 3\sigma$  from the input fiducial value ( $-0.35$ ) while  $M_{turn}$  is  $\sim 2\sigma$  below the fiducial value. Therefore, even for relatively modest differences in the HMF, the inferred astrophysical parameters can be strongly biased.

## 5 JOINTLY RECOVERING THE HMF AND ASTROPHYSICAL PARAMETERS

Thus far, we have explored the relative biases in the inferred model parameters owing to incorrectly assuming the HMF for our underlying astrophysical model. Generally speaking, this assumption of a fixed HMF stems from the reduced uncertainty in the HMF models relative to those describing the astrophysical properties of the galaxies allowing us to limit the computational complexity for our inference pipelines. In this section, we explore the consequences of relaxing this assumption and attempt to jointly constrain both the astrophysical parameters as well as the underlying model for the HMF.<sup>11</sup> For this, we assume the same mock 21-cm PS observation assuming a ST HMF.

### 5.1 Generalized HMF

Throughout the literature, there are a large number of different analytic fitting functions that have been derived to fit the HMFs of  $N$ -body simulations (e.g. Jenkins et al. 2001; Warren et al. 2006; Reed et al. 2007; Tinker et al. 2008; Tinker et al. 2010; Angulo et al. 2012; Watson et al. 2013; Diemer 2020). These can sometimes contain several parameters which may or may not have complex scalings with redshift and/or cosmological parameters. Deriving a universal, generalized HMF to describe all  $N$ -body results is beyond the scope of this work; however, we can draw from the more common parametrizations in the literature to provide a simple, basic example.

The HMF can simply be defined as

$$\frac{dn}{dM} = f(\sigma) \frac{\bar{\rho}_m}{M} \frac{d \ln \sigma^{-1}}{dM}, \quad (24)$$

where  $\sigma$  is the root mean square linear overdensity of the density field smoothed by a top-hat filter of radius,  $R$ , containing a mass  $M$  at the mean cosmic matter density,  $\bar{\rho}_m$ . Defined in this way, the function  $f(\sigma)$  encapsulates the intrinsic behaviour for any assumed model of the HMF.

Since our mock observation is generated assuming a ST HMF, we first consider a generalized five parameter model of the following form:

$$f(\sigma) = a [(b\nu)^c + 1] \nu^e \exp(-d\nu^2), \quad (25)$$

where  $\nu = \delta_c/\sigma$  and  $\delta_c$  is the critical density. Defined in this way, we have five free parameters which we simply denote  $a$ ,  $b$ ,  $c$ ,  $d$ , and  $e$ , which is more flexible than the actual ST HMF. As a result, this form slightly differs from the actual ST HMF, whereby their HMF is parametrized assuming a  $\nu' = \alpha\delta_c/\sigma$  in place of our  $\nu$ , which results in slightly different amplitudes for the various coefficients. Nevertheless, adopting  $a = 0.2406$ ,  $b = 0.85$ ,  $c = -0.35$ ,  $d = 0.365$ , and  $e = 1.0$  can mimic the behaviour for the ST HMF model. Generally speaking, this functional form can recover most HMFs in the literature, with the largest discrepancies being those models with coefficients with redshift dependence (e.g. Reed et al. 2007; Tinker et al. 2008; Tinker et al. 2010; Watson et al. 2013).

While this five parameter HMF should be able to recover the ST HMF, it will also be illustrative to explore the relative impact of the assumed functional form of the HMF on our inferred parameters.

<sup>11</sup>Alternatively, rather than expanding the parameter set to also include the HMF parameters, if one was only interested in the astrophysical parameters one could instead consider performing Bayesian model averaging of the recovered posteriors obtained from assuming a broad range of HMFs from the literature.

That is, if we assume an incorrect function form incapable of mimicking the true form of the HMF, how significant will the relative biases be? Therefore, we shall consider a secondary model whereby we set  $e = 0$  in equation (25), thus considering a simpler four parameter generalized HMF. This choice is somewhat arbitrary as we could theoretically remove any model parameter, however, the removal of the  $\nu$  multiplicative term should have a more notable impact on the HMF shape and amplitude and correspondingly the astrophysical parameters. In future, one could envisage using Bayesian evidence to more robustly determine better generalized functional forms for characterizing the HMF.

### 5.2 Mock observation using only the 21-cm PS

Since we have expanded the dimensionality of our model to 13 individual parameters, we are required to generate a new data base of models. Specifically, we generated a new data base of 200 000 models for both the four and five parameter generalized form of the HMF. For the five new HMF parameters we adopt relatively broad priors which are summarized in Table 3. In Fig. 8, we demonstrate our marginalized posteriors for our model parameters using SWYFT. Along the diagonal we show the marginalized 1D posteriors whereas below the diagonal we provide the 95th percentile joint marginalized posteriors. In each, the grey (orange) curves correspond to the recovered posteriors assuming the five (four) parameter generalized form for the HMF. For reference, we equally show the marginalized posteriors when assuming the correct ST HMF (black curves, Fig. 5). Finally, in Table 3 we summarize the marginalized constraints and 68th percentile uncertainties for each of the 13 model parameters.

As one would expect, by simultaneously constraining the astrophysical and HMF parameters, we now recover considerably broader posteriors along with several additional strong model parameter degeneracies. The most significant of these degeneracies is between the various normalization quantities, e.g.  $\log_{10}(f_{*,10})$ ,  $\log_{10}(f_{\text{esc},10})$  and now the HMF normalization,  $a$ . As these are multiplicative within our 21-cm model, the enlarged posterior volumes owing to these degeneracies play a role in broadening the overall posteriors of all the model parameters. Realistically, rather than attempting to constrain all three normalization quantities independently, one could parametrize the product of  $\log_{10}(a \times f_{*,10})$  to minimize the relative strength of these degeneracies, keeping  $f_{\text{esc}}$  independent as the  $f_* - f_{\text{esc}}$  degeneracy can be reduced following the inclusion of UV LFs. Nevertheless, as a first step we keep these three parameters independent to explore the relative amplitude of these uncertainties when attempting to constrain all model parameters.

Focusing first on the five parameter HMF model, we can correctly recover almost all of our astrophysical parameters to within the much broader 68th percentile uncertainties, with  $f_{*,10}$  being consistent within the 95th percentile uncertainties ( $M_{\text{turn}}$  is right at the edge of the 68th percentile). The significant deviation in  $f_{*,10}$  comes from the noted degeneracy between the HMF normalization,  $a$  and  $\log_{10}(f_{*,10})$  and  $\log_{10}(f_{\text{esc},10})$ . For example, our recovered constraint on  $\log_{10}(a)$  is  $-0.16$  relative to the fiducial value of  $-0.62$  ( $\sim 1.5\sigma$  larger). To compensate for the larger  $a$ , we must significantly reduce  $\log_{10}(f_{*,10})$  ( $-1.83$  instead of our assumed  $-1.1$ , or approximately  $\sim 1.5\sigma$  lower). The larger than expected  $\log_{10}(f_{\text{esc},10})$  corresponds to the fact that we have a smaller stellar component, requiring a higher escape fraction to ensure reionization occurs at the correct point in time. For the HMF parameters, other than the strongly offset  $a$ , the remaining parameters are within their 68th percentiles, owing to the fact that the relative uncertainties are quite large. Interestingly, the



**Table 3.** A summary of the recovered parameter constraints plus 68th percentiles when we additionally vary the underlying HMF along with the astrophysical parameters within our inference pipeline. For this, we consider the same mock observation of the 21-cm PS, assuming foreground wedge avoidance and a 1000 h observation with the SKA and a ST HMF (Sheth et al. 2001). Additionally, we provide the recovered constraints when also including the additional constraining power from UV LFs (see the text for further details).

Astrophysical parameters	$\log_{10}(f_{*,10})$	$\alpha_*$	$\log_{10}(f_{\text{esc},10})$	$\alpha_{\text{esc}}$	$\log_{10}(M_{\text{turn}})$ ( $M_{\odot}$ )	$t_*$	$\log_{10}\left(\frac{L_{X<2\text{keV}}}{\text{SFR}}\right)$ ( $\text{erg s}^{-1} M_{\odot}^{-1} \text{yr}^{-1}$ )	$E_0$ (keV)
Mock observation	-1.1	0.5	-1.30	-0.35	8.55	0.5	40.5	0.5
21-cm PS only								
ST HMF	$-1.14^{+0.10}_{-0.13}$	$0.48^{+0.13}_{-0.19}$	$-1.25^{+0.12}_{-0.14}$	$-0.29^{+0.14}_{-0.08}$	$8.60^{+0.27}_{-0.15}$	$0.48^{+0.14}_{-0.15}$	$40.47^{+0.03}_{-0.03}$	$0.49^{+0.03}_{-0.02}$
Five-parameter HMF	$-1.83^{+0.48}_{-0.47}$	$0.66^{+0.29}_{-0.24}$	$-1.25^{+0.30}_{-0.23}$	$-0.57^{+0.30}_{-0.21}$	$8.82^{+0.31}_{-0.27}$	$0.64^{+0.32}_{-0.16}$	$40.54^{+0.06}_{-0.05}$	$0.53^{+0.08}_{-0.05}$
Four-parameter HMF	$-1.47^{+0.48}_{-0.43}$	$0.30^{+0.36}_{-0.29}$	$-1.23^{+0.23}_{-0.24}$	$-0.37^{+0.28}_{-0.20}$	$9.08^{+0.28}_{-0.28}$	$0.56^{+0.28}_{-0.18}$	$40.50^{+0.06}_{-0.05}$	$0.54^{+0.07}_{-0.07}$
Five-parameter (joint norm)	-	$0.72^{+0.25}_{-0.23}$	$-1.23^{+0.28}_{-0.27}$	$-0.65^{+0.26}_{-0.22}$	$8.68^{+0.30}_{-0.29}$	$0.71^{+0.25}_{-0.30}$	$40.54^{+0.06}_{-0.06}$	$0.54^{+0.05}_{-0.05}$
Four-parameter (joint norm)	-	$0.27^{+0.36}_{-0.22}$	$-1.26^{+0.25}_{-0.24}$	$-0.39^{+0.19}_{-0.27}$	$9.24^{+0.21}_{-0.30}$	$0.41^{+0.40}_{-0.12}$	$40.55^{+0.06}_{-0.06}$	$0.52^{+0.05}_{-0.04}$
21-cm PS + UV LFs								
ST HMF	$-1.11^{+0.13}_{-0.13}$	$0.45^{+0.08}_{-0.09}$	$-1.20^{+0.13}_{-0.11}$	$-0.30^{+0.05}_{-0.05}$	$8.47^{+0.16}_{-0.12}$	$0.58^{+0.16}_{-0.12}$	$40.49^{+0.03}_{-0.02}$	$0.49^{+0.04}_{-0.03}$
Five-parameter HMF	$-1.68^{+0.41}_{-0.50}$	$0.57^{+0.11}_{-0.09}$	$-1.06^{+0.22}_{-0.17}$	$-0.48^{+0.12}_{-0.09}$	$8.60^{+0.36}_{-0.27}$	$0.59^{+0.25}_{-0.22}$	$40.53^{+0.06}_{-0.05}$	$0.55^{+0.07}_{-0.06}$
Four-parameter HMF	$-1.42^{+0.56}_{-0.50}$	$0.68^{+0.12}_{-0.13}$	$-1.04^{+0.22}_{-0.18}$	$-0.55^{+0.15}_{-0.19}$	$8.88^{+0.24}_{-0.25}$	$0.50^{+0.20}_{-0.22}$	$40.56^{+0.06}_{-0.08}$	$0.53^{+0.06}_{-0.04}$
Five-parameter (joint norm)	-	$0.55^{+0.18}_{-0.13}$	$-1.27^{+0.24}_{-0.24}$	$-0.40^{+0.14}_{-0.12}$	$8.77^{+0.33}_{-0.24}$	$0.54^{+0.22}_{-0.20}$	$40.50^{+0.07}_{-0.06}$	$0.54^{+0.08}_{-0.06}$
Four-parameter (joint norm)	-	$0.33^{+0.14}_{-0.09}$	$-1.01^{+0.19}_{-0.24}$	$-0.44^{+0.15}_{-0.21}$	$9.08^{+0.23}_{-0.27}$	$0.46^{+0.28}_{-0.18}$	$40.54^{+0.06}_{-0.06}$	$0.52^{+0.06}_{-0.05}$
HMF parameters	-	$\log_{10}(a \times f_{*,10})$	$\log_{10}(a)$	$\log_{10}(b)$	c	$\log_{10}(d)$	e	-
Mock observation	-	-1.72	-0.62	-0.07	-0.35	-0.44	1.0	-
Prior ranges	-	[-5.00, 1.70]	[-2.00, 1.70]	[-2.00, 1.70]	[-3.50, 3.50]	[-1.00, 0.00]	[-2.50, 2.50]	-
21-cm PS + UV LFs								
Five-parameter HMF	-	-	$-0.16^{+0.36}_{-0.33}$	$0.24^{+0.44}_{-0.37}$	$0.50^{+0.84}_{-0.84}$	$-0.47^{+0.07}_{-0.06}$	$0.57^{+0.85}_{-0.87}$	-
Four-parameter HMF	-	-	$0.15^{+0.23}_{-0.39}$	$0.42^{+0.22}_{-0.31}$	$-0.43^{+1.03}_{-0.63}$	$-0.52^{+0.06}_{-0.05}$	-	-
Five-parameter (joint norm)	-	$-1.69^{+0.36}_{-0.39}$	-	$0.18^{+0.49}_{-0.65}$	$0.76^{+0.70}_{-0.93}$	$-0.47^{+0.05}_{-0.08}$	$0.21^{+0.70}_{-0.88}$	-
Four-parameter (joint norm)	-	$-1.45^{+0.27}_{-0.29}$	-	$0.33^{+0.20}_{-0.26}$	$0.46^{+0.56}_{-1.17}$	$-0.54^{+0.05}_{-0.06}$	-	-
21-cm PS + UV LFs								
Five-parameter HMF	-	-	$-0.27^{+0.23}_{-0.35}$	$0.02^{+0.57}_{-0.58}$	$0.13^{+0.79}_{-0.86}$	$-0.51^{+0.06}_{-0.05}$	$0.27^{+0.82}_{-0.97}$	-
Four-parameter HMF	-	-	$-0.29^{+0.28}_{-0.30}$	$0.31^{+0.26}_{-0.70}$	$-0.43^{+1.46}_{-0.68}$	$-0.59^{+0.07}_{-0.06}$	-	-
Five-parameter (joint norm)	-	$-1.82^{+0.35}_{-0.31}$	-	$0.13^{+0.49}_{-0.63}$	$0.15^{+0.77}_{-0.68}$	$-0.49^{+0.06}_{-0.05}$	$0.86^{+0.58}_{-1.15}$	-
Four-parameter (joint norm)	-	$-1.69^{+0.31}_{-0.24}$	-	$0.35^{+0.23}_{-1.21}$	$-0.62^{+1.07}_{-0.68}$	$-0.58^{+0.05}_{-0.05}$	-	-

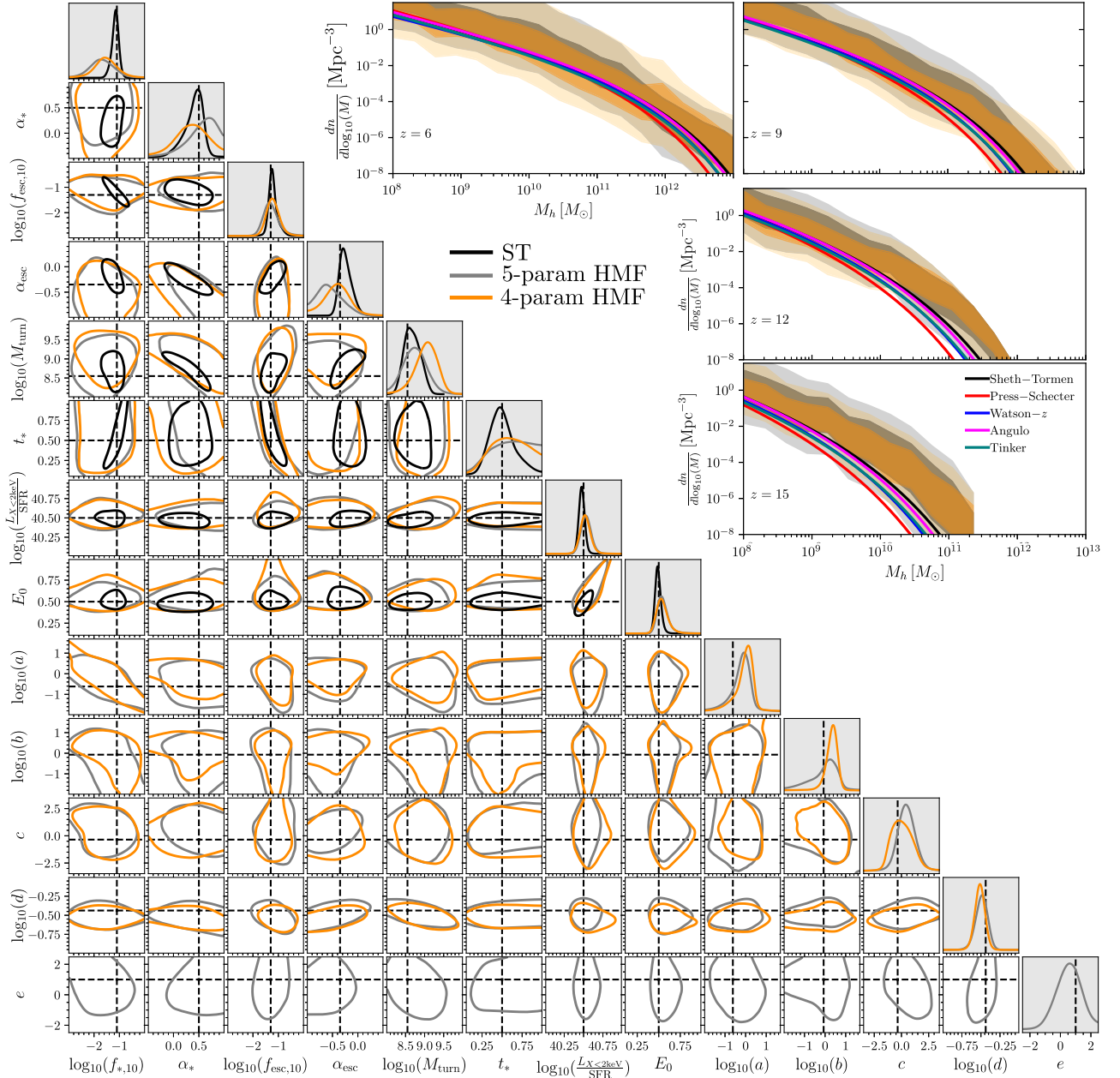
parameter  $d$  is quite strongly constrained compared to all other HMF parameters, which corresponds to the location and amplitude of the exponential drop-off within the HMF. Amongst the various HMF parameters, there are no significantly strong degeneracies other than that between  $b$  and  $c$ .

For the four parameter HMF, we recover relatively comparable constraints on our astrophysical parameters as the five parameter model. The largest differences occur for  $M_{\text{turn}}$ ,  $\alpha_*$ , and  $\alpha_{\text{esc}}$ . For  $\alpha_*$  and  $\alpha_{\text{esc}}$  we are simply on the opposite sides of the strong  $\alpha_* - \alpha_{\text{esc}}$  degeneracy relative to the five parameter HMF model. However, for the four parameter model we prefer a much higher  $M_{\text{turn}}$  ( $\sim 9.08$ ) which is due to the strong degeneracy between  $M_{\text{turn}}$  and the exponential term in the HMF,  $d$ . As this term,  $d$ , is relatively strongly constrained and degenerate with  $M_{\text{turn}}$ , its offset from the expected value leads to the observed increase in  $M_{\text{turn}}$ .

Overall, for the astrophysical parameters the relative amplitude of the uncertainties increase due to the simultaneous constraining of the HMF parameters by a factor of  $\sim 4$  for  $\log_{10}(f_{*,10})$  and  $\sim 2-3$  for the remainder. Relatively speaking for all parameters in our expanded model, the four parameter HMF model has slightly reduced marginalized uncertainties owing to the fact that we have one less parameter in our model. As a consequence, this will marginally amplify any respective biases if the assumed form of the HMF is unable to match that of the input ST HMF. Indeed, for the four parameter HMF model we recover a larger than expected value for  $M_{\text{turn}}$  (over  $\sim 2\sigma$ ) to compensate for the bias in  $d$ .

In order to better understand the sensitivity of the model parameters to the exponential term,  $d$ , and the preference for a decreasing  $d$  between the four and five parameter HMF models, we explore the sensitivity of this parameter to the various summary statistics of the 21-cm signal. For this, we consider the same astrophysical and HMF model parameter values as the mock observation with the ST HMF (except with  $e = 0$  for the four parameter HMF model). In Fig. 9, we consider: (i) the reionization history, (ii) the global brightness temperature, (iii) the 21-cm PS and (iv) the corresponding HMFs. In all, the black (red) curves correspond to the five (four) parameter HMF model. For this four parameter HMF, we then vary  $\log_{10}(d)$  in increments of  $-0.05$  from  $-0.45$  to  $-0.65$ , as highlighted by the light-blue curves.

Quite simply, by removing the fifth parameter ( $e = 0$ ) while keeping all others parameter values fixed results in a considerably different behaviour for the 21-cm signal. By decreasing the value of  $d$  (i.e. softening the exponential cut-off of the HMF) from the expected value in the four-parameter model, we can quickly recover statistics of the 21-cm signal consistent with the five-parameter model. However, modifications to  $d$  alone cannot do all the work, and we further require other shifts in the model parameters in order to match the amplitude and shape of the 21-cm PS. Decreasing  $d$  leads to an increasing relative amplitude for the HMF which increases the number of sources (e.g. UV ionizing and X-ray photons) causing reionization and other 21-cm milestones to occur earlier. Therefore, owing to the sensitivity of the HMF to  $d$ , it is both the most strongly

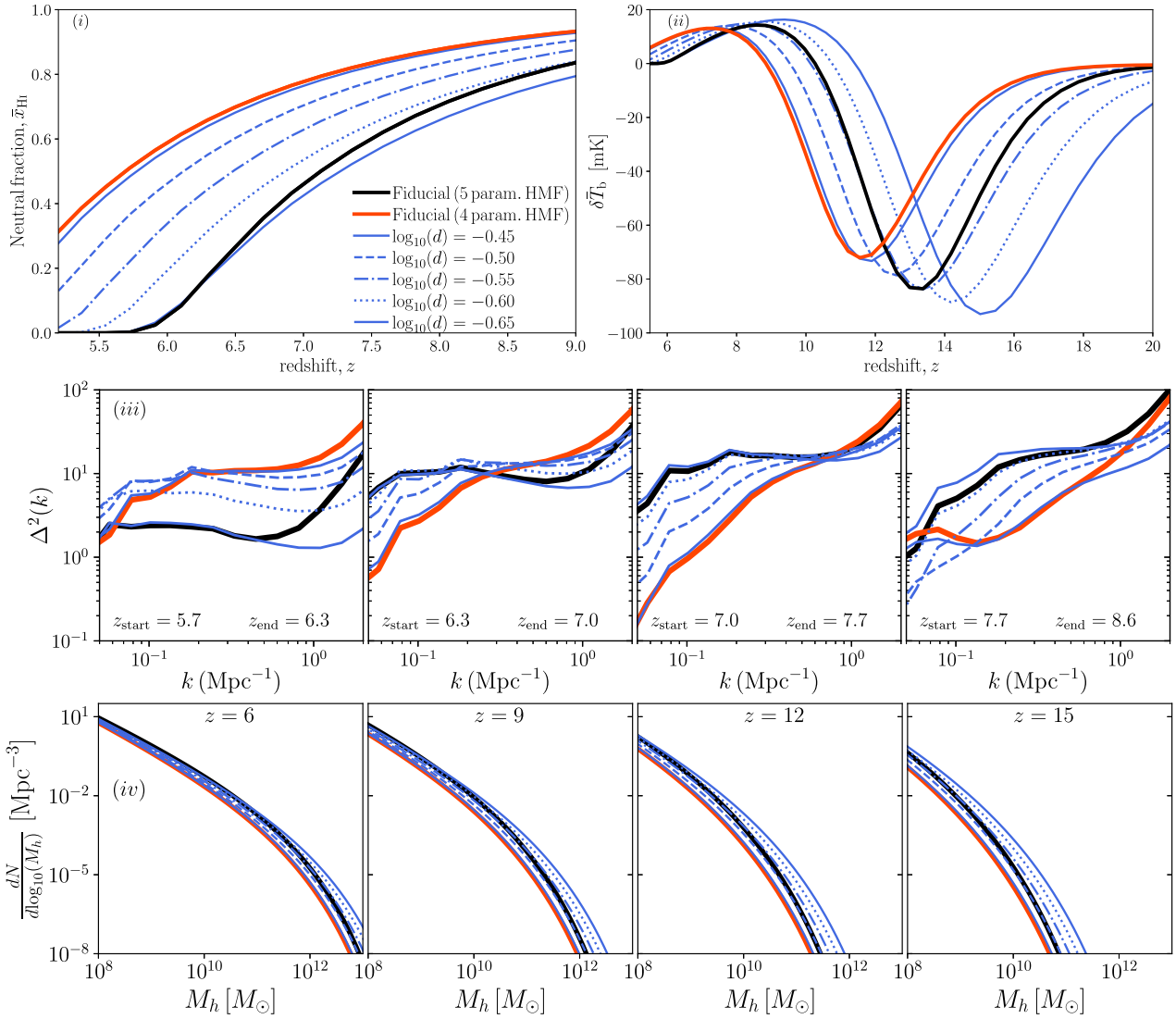


**Figure 8.** The recovered one and 2D marginalized posteriors on our joint astrophysical and general HMF function parameters assuming a mock 1000 h observation of the 21-cm PS assuming foreground avoidance with the SKA. Note, our mock observation assumes the Sheth–Tormen HMF. The grey (orange) contours correspond to a five (four) parameter HMF (see the text for further details). The vertical and horizontal black dashed lines denote our fiducial astrophysical parameter set and the corresponding values for the Sheth–Tormen HMF. Above the diagonal, we provide the marginalized posteriors on the HMF when simultaneously constraining both the astrophysical and HMF parameters assuming our same mock observation. Each panel corresponds to  $z = 6, 9, 12$  and  $15$ , respectively. Within each panel we show the 68th (dark) and 95th (light) marginalized posteriors on the recovered HMF. The grey (orange) contours correspond to the marginalized posteriors assuming our five (four) parameter HMF. For reference, we overlay several different HMFs as explored earlier. The vertical offset between the marginalized posteriors and the analytic expressions is due to the degeneracy between the HMF normalization,  $a$ , and the normalizations of  $f_*$  and  $f_{\text{esc}}$  (see the text for further details).

constrained and also most susceptible to being offset from its true value. The behaviour observed here is consistent with the trends in the biased parameters observed for the four-parameter model found above. The lower value of  $d$  increases the amplitude in the HMF allowing for the EoR to occur earlier; however, in order not to overproduce the too many haloes, we must subsequently shift to a higher  $M_{\text{turn}}$ .

### 5.3 Differentiating between HMFs

As established earlier, fixing the underlying HMF can lead to the biased recovery of the astrophysical parameters. Alternatively, one can attempt to simultaneously constrain the astrophysical parameters and a generalized form for the HMF in order to mitigate the potential biases, at the cost of considerably weakened constraints (factor of

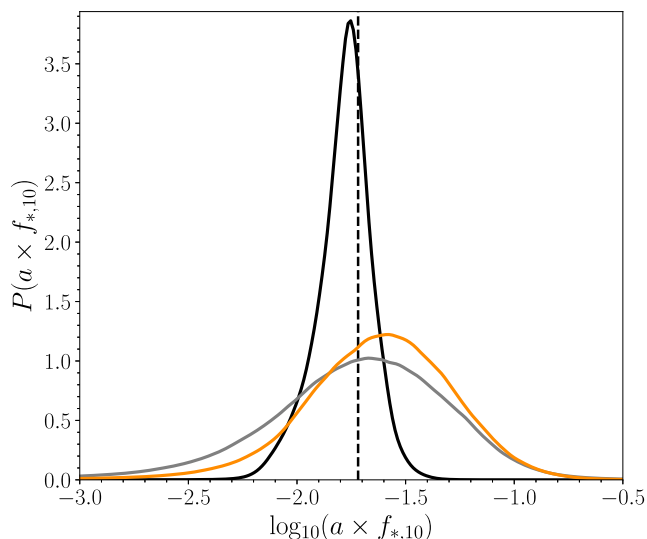


**Figure 9.** The impact of changing the HMF constant which controls the exponential cut-off,  $d$ , in the four parameter HMF parametrization. We compare the behavior of this parameter relative to the fiducial astrophysical model with a Sheth–Tormen HMF (black curve; the five parameter HMF model designed to mimic the ST HMF) and also a four parameter HMF parametrization (red curve) with one HMF parameter removed (the same parameters as the five parameter HMF model to match the ST HMF except with  $e = 0$ ). Panel (i) corresponds to the reionization history, (ii) corresponds to mean 21-cm brightness temperature signal, (iii) shows the 21-cm PS for a couple of redshifts, and (iv) the HMF. We consider  $\log_{10}(d) = (-0.45, -0.5, -0.55, -0.6, -0.65)$ , demonstrating the sensitivity of the 21-cm signal to this parameter.

$\sim 2 - 4$ ). Although the marginalized posteriors on the individual HMF parameters are rather broad, the important question is whether or not these translate to marginalized posteriors on the HMF as a whole which are capable of distinguishing between various HMFs in the literature.

Again, to obtain our posteriors on the HMFs given our data base of simulated models, we are required to retrain SWYFT (equivalent to extracting the SFRD posteriors in Section 4.2). Strictly speaking, we train a new set of networks for each HMF at each redshift. In the upper-right corner of Fig. 8, we present the marginalized 68th (95th) posteriors for the HMF obtained from the same mock 21-cm PS observation at  $z = 6, 9, 12$  and  $15$ . The grey (orange) curves correspond to the five (four) parameter model for the HMF, respectively. For reference, we additionally overlay the five HMFs considered in our earlier analysis, namely ST (black), PS (red), Watson- $z$  (blue), Angulo (magenta), and Tinker (teal).

Immediately evident from Fig. 8 is that the marginalized posteriors on the HMF are broader than the difference between the various analytic HMFs considered in this work. Therefore, simultaneously constraining the astrophysical and HMF model parameters from just a 21-cm PS observation with the SKA will not be sufficient to clearly distinguish between the various HMF models available in the literature. However, despite the relatively broad posteriors on the individual HMF model parameters, the marginalized posteriors on the actual HMF are fairly tightly constrained around the input ST HMF. Note, in both cases, we find the recovered HMF posteriors marginally overestimate the true underlying HMF, with the level of discrepancy increasing for higher redshifts. However, this is not unexpected given that we prefer a lower  $d$  in both cases, which causes an increase to the amplitude and slope of the HMF for increasing redshift (see the bottom panel of Fig. 9). Further, as established earlier we also prefer an increase in the overall HMF normalization,  $a$ , at the



**Figure 10.** The recovered marginalized 1D posteriors on the joint normalization,  $a \times f_{*,10}$ , obtained after retraining SWYFT for include this normalization product in our joint astrophysical and general HMF model assuming a mock 1000 h observation of the 21-cm PS assuming foreground avoidance with the SKA. The grey (orange) curves correspond to our five (four) parameter HMF (see the text for further details). The vertical black dashed line denote our fiducial value for this joint normalization and the black curves are the posteriors from Fig. 5 when assuming the correct ST HMF multiplied by the known ST HMF normalization. Note, we only show the 1D marginalized posteriors for  $a \times f_{*,10}$  as the 1D and 2D marginalized posteriors for all remaining model parameters should remain mostly unchanged (modulo SBI stochasticity) as those in Fig. 8 as the underlying training set remains unchanged.

cost of a decreased  $f_*$  and  $f_{\text{esc}}$  which further drives up the recovered HMF posteriors. Nevertheless, this demonstrates that simultaneously inferring the astrophysical and HMF model parameters from the 21-cm PS alone does a reasonable job at constraining the underlying HMF.

#### 5.4 Reducing the normalization parameters

Earlier, we found that allowing  $a$ ,  $f_{*,10}$ , and  $f_{\text{esc},10}$  to be independently constrained in our model resulted in notably large offsets in their recovered values. Therefore now, we repeat our analysis, reducing these three parameters down to two, the product  $a \times f_{*,10}$  and  $f_{\text{esc},10}$ . Within 21CMFAST  $f_{*,10}$  is always multiplicative with the HMF, as it is used to determine star formation, either in the production of UV or X-ray photons, thus these two quantities will always be correlated.  $f_{\text{esc},10}$  on the other hand only becomes important during the EoR, thus it can act somewhat independently.

In Fig. 10, we provide the 1D marginalized posteriors for the joint normalization,  $a \times f_{*,10}$ , for the same mock observation of the 21-cm PS with the SKA after retraining SWYFT for our updated joint astrophysical and HMF model. Again, the black curve corresponds to the assuming the correct ST HMF (black curves, Fig. 5) after multiplying the recovered  $f_{*,10}$  by the known ST HMF normalization) whereas the black dashed vertical line is the fiducial joint normalization. The grey (orange) curves correspond to the 5 (4) parameter HMF model. Note, here we only show the 1D marginalized posteriors for  $a \times f_{*,10}$  as the 1D and 2D marginalized posteriors for all remaining parameters should remain the same as those in Fig. 8, as the underlying training set remains unchanged.

However, in practice owing to inherent stochasticity within the SBI framework, some minor differences are recovered. Nevertheless, in Table 3 we summarize the 68th percent marginalized uncertainties for this full model parameter set with joint normalization. After considering,  $a \times f_{*,10}$ , we now recover unbiased constraints on this joint normalization, centred around the fiducial value. This is a notable improvement over the large offsets recovered when considered these parameters individually (Fig. 8). Therefore, as expected, considering this normalization product has improved the visual performance of our overall model.

#### 5.5 Mock 21-cm PS plus UV LF observation

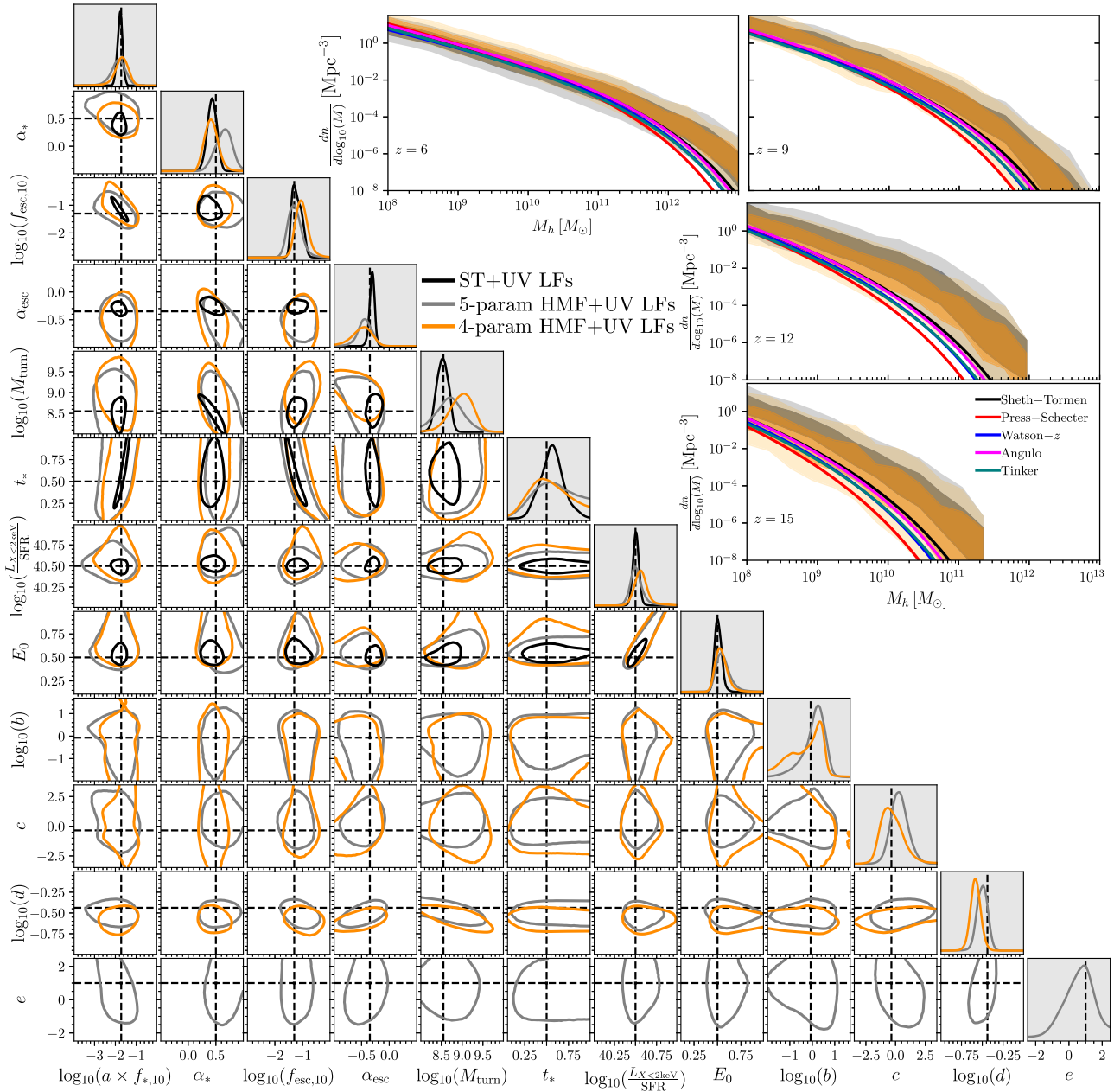
Finally, we consider the impact on the recovered model parameters following the inclusion of UV LF data to our mock observation. In Fig. 11, we provide the marginalized 1D and 2D posteriors for our mock 21-cm PS observation plus the same UV LFs as considered previously (namely at  $z = 6, 7, 8$  and 10). Specifically, we only present the results when considering the normalization product,  $a \times f_{*,10}$ . However, for completeness we include both this joint normalization and when keeping  $a$  and  $f_{*,10}$  independent in Table 3. Finally, in the upper right corner, we also provide the marginalized 68th and 95th percentiles on the resultant HMF following the joint recovery of the astrophysical and HMF model parameters.

Generally speaking, following the inclusion of the UV LFs into our mock observation, the relative improvements in our constraints are more modest compared to a model considering only the astrophysical parameters (i.e. assuming a fixed choice of HMF). However, this is driven by the increased model complexity in trying to simultaneously recover both the astrophysical and HMF parameters. Nevertheless, including the UV LFs improve the overall constraints primarily on  $\alpha_{\text{esc}}$  and  $a \times f_{*,10}$ , which subsequently improves the constraints on  $\alpha_{\text{esc}}$  and  $f_{\text{esc},10}$  by breaking the degeneracy. For example, in the absence of the UV LFs, we recovered a relatively larger and more skewed  $\alpha_*$  (0.72) for the five-parameter HMF, however, using the UV LF data, this almost completely removes this offset with  $\alpha_*$  now  $\sim 0.55$  compared to the assumed 0.50.

Both the four- and five-parameter HMF models still exhibit offsets in  $\log_{10}(d)$  as we have encountered previously. For example, for the five-parameter model we recover  $-0.49 \pm 0.06$  which remains borderline consistent with the 68th percentile uncertainties to the fiducial value of  $-0.44$ . However, the four-parameter HMF model is now inconsistent at  $\sim 3\sigma$ , preferring  $-0.59 \pm 0.05$ .

These lower than expected values of  $\log_{10}(d)$  affect our ability to constrain all other remaining model parameters, for example pushing  $M_{\text{turn}}$  to larger values to compensate for the increasing HMF function amplitude. In turn, this leads to an increasing  $f_{\text{esc}}$  to allow more UV photons to escape to account for the fact we have fewer lower mass galaxies (increased  $M_{\text{turn}}$ ). For the five parameter HMF, owing to the more modest offset this manifests as a lower  $a \times f_{*,10}$  and also a lower  $\alpha_{\text{esc}}$  which increases the escape fraction of the lower mass galaxies. On the other hand for the four parameter HMF, we need a significant increase in  $f_{\text{esc},10}$  to compensate for the much larger offset in  $M_{\text{turn}}$  and  $d$ . Therefore, removing a model parameter from our generalized HMF (i.e. the parameter  $e$ ) results in a less flexible function form for the HMF leading to stronger biases in the inferred astrophysical parameters. As a result, if one attempts to simultaneously recover the astrophysical and HMF model parameters, the assumed functional form for the HMF must be sufficiently flexible.

Following the inclusion of the UV LFs, focusing on the five parameter model which matches the functional form of the assumed



**Figure 11.** The same as Fig. 10 except also including the UV LFs at  $z = 6, 7, 8$  and  $10$  as observational priors.

ST HMF, jointly recovering the astrophysical and HMF model parameters results in fractional uncertainties on our astrophysical parameters of between a factor of  $\sim 2 - 3 \times$  larger than assuming a fixed HMF. Importantly, the recovered marginalized posteriors from simultaneously fitting the HMF jointly with the astrophysical parameters encapsulate the posteriors when fixing the HMF to the correct ST HMF. Further, they recover the same relative degeneracies as evidenced by the similar posterior contours. These more broadened posteriors additionally account for the relative biases in the astrophysical parameters when we previously incorrectly assumed the form of the HMF. That is, the biased posteriors recovered in Fig. 7 are contained within these broadened posteriors. Except for the PS HMF, which clearly prefers lower than allowed values of  $M_{\text{turn}}$ . Therefore, by relaxing the assumption of choosing a fixed HMF, we

are able to obtain more conservative, unbiased constraints on our astrophysical parameters.

In principle, one could also consider other observational constraints on our joint model parameters, which could further improve our overall constraining power limiting the impact of increasing the model complexity to jointly recover the HMF. Nevertheless, these relative increases in the marginalized uncertainties, when jointly recovering the HMF, are not prohibitive when it comes to interpreting real-world data. In fact, in effectively being agnostic to the underlying HMF, the precision of these recovered constraints are more conservative with respect to our ignorance of our modelling uncertainties. However, further work is required to determine the preference for a lower than expected value for the exponential cut-off parameter,  $d$ , and ways to mitigate/eliminate it.

Focusing on the recovered HMF posteriors, we see an almost factor of three reduction in the posterior width at  $z = 6$  with these posteriors tightly bounding the various HMFs in the literature. However, the relative improvements decrease for increasing redshift, highlighting that the vast majority of the constraining power comes from the UV LFs at  $z = 6$ . We do observe a more modest shrinking of the posteriors at  $z = 9$ , where we still have observational data (at  $z = 10$ ), but no relative improvement at  $z = 12$  or  $z = 15$ . Therefore, to be able to distinguish between the various HMFs in the literature, we would require improved UV LF constraints spanning out to higher redshifts. Already, recent *JWST* results are beginning to place limits on the UV LFs extending to redshifts higher than considered here (e.g. Castellano et al. 2022; Harikane et al. 2022; Naidu et al. 2022; Atek et al. 2023; Bouwens et al. 2023; Donnan et al. 2023; Labbé et al. 2023; Willott et al. 2024). Further, relaxing our assumption of only considering UV galaxies at  $M_{UV} < -20$  allowing us to access the bright end galaxies might further help improve our ability to constrain the underlying HMF from our mock 21-cm PS observation with the SKA by constraining the higher mass end. We shall look into the improvements available when considering both of these in future work.

## 6 CONCLUSIONS

We are rapidly approaching the first detection of the 21-cm signal during reionization. To gain insights into the properties of the first generation of galaxies responsible for driving reionization we are required to develop Bayesian inference pipelines to interpret the observation data. Crucially, when developing these inference frameworks we are required to make various assumptions and simplifications in our theoretical models in order to make them computationally efficient enough to be used for inference. However, it is crucial that we understand the consequences of these choices. One such assumption that is typically adopted is that of a fixed HMF. Which implies that our inferred model parameter constraints are tied to our assumed choice of HMF.

Within this work, we explored the consequences of incorrectly assuming the HMF on the inferred astrophysical parameters assuming a mock observation of the 21-cm power spectrum (PS). Specifically, we consider a 1000 h observation of the 21-cm PS assuming foreground avoidance with the forthcoming SKA. Throughout we modelled the 21-cm signal using 21CMFAST (Mesinger & Furlanetto 2007; Mesinger et al. 2011; Murray et al. 2020) adopting the flexible UV galaxy model parametrization of Park et al. (2019). This work extends on earlier work by Mirocha et al. (2021) who explored the impact of assuming a HMF along with the stellar population synthesis model on the global 21-cm signal. However, this used a simplified model for simulating the 21-cm signal, whereas here we perform full 3D reionization simulations in order to study the 21-cm PS. To recover our inferred parameter constraints we performed SBI using the publicly available software, SWYFT (Miller et al. 2021) which performs MNRE to learn the likelihood-to-evidence ratio of our simulated data to obtain the marginalized posteriors on our astrophysical parameters.

Adopting the ST HMF (Sheth et al. 2001) as our fiducial model HMF, we then recovered our model astrophysical parameters assuming five difference choices for the underlying HMF in order to encapsulate the breadth of different models in the literature. In addition to the ST HMF, we also considered the PS (Press & Schechter 1974), Angulo (Angulo et al. 2012), Tinker (Tinker et al. 2010) and the redshift evolving Watson HMF (Watson et al. 2013). When considering just the 21-cm PS alone, we encountered strong

biases of up to  $\sim 3-4\sigma$  on our recovered astrophysical parameters, most notably in the mass dependence of the escape fraction,  $\alpha_{\text{esc}}$  and  $M_{\text{turn}}$  the characteristic turn-over mass describing the typical masses of star-forming galaxies.

The source of these biases are driven by the underlying model seeking to match the SFRD over the entire cosmic history. The differences in the HMF lead to the cosmic 21-cm milestones (reionization, X-ray heating or Ly  $\alpha$  coupling) happening earlier or later depending on if they over or under predict the number density of sources. Thus, the biases in astrophysical parameters compensate for the HMF differences to ensure the cosmic SFRD correctly matches that of the mock observation.

Next, we folded in observed UV LFs with our 21-cm PS observation in order to improve our constraints on our model parameters. Primarily, the inclusion of the UV LFs constrains the stellar content of the UV galaxies strongly constraining  $f_*$ , both through its normalization and mass dependence. This in turn breaks the corresponding degeneracy with the escape fraction,  $f_{\text{esc}}$ . In doing so, this restricts the resultant posterior volume available to match the true SFRD and compensate for the incorrect HMF choice. Primarily, this drives biases in  $M_{\text{turn}}$  and  $\alpha_{\text{esc}}$ . Although the relative amplitude of these biases are reduced compared to just the 21-cm PS, the reduced marginalized uncertainties still lead to significant biases of  $\sim 3\sigma$  for both parameters.

Given that incorrectly assuming the HMF can lead to large biases in our astrophysical parameters, next we explored whether we could jointly recover the astrophysical parameters along with the underlying HMF from the same mock 21-cm PS observation with the SKA. To do so, we constructed a simple five parameter HMF model and simultaneously recovered these parameters in addition to our eight astrophysical parameters. Further, we also considered the case where we removed one of these five HMF model parameters, in order to explore the consequences of jointly fitting a HMF that is inflexible to the true underlying ST HMF.

By constraining this extended model with a generalized form for the HMF, our astrophysical parameters are recovered with a factor of  $\sim 2-4$  larger marginalized uncertainties, owing to the increased model complexity. However, in doing so our constraints are now unbiased by our choice of HMF and are thus more conservative. Note, we still recovered small offsets in some of our astrophysical parameters, namely  $\alpha_*$  and  $\alpha_{\text{esc}}$ , but this is driven by a corresponding offset in the opposite direction for one of the HMF model parameters,  $d$  which controls the exponential cut-off. In the case of our four parameter model, which is designed to be inflexible, we recovered considerably stronger skews in  $M_{\text{turn}}$  and  $d$ . Therefore, by assuming a generalized HMF incapable of matching the observation data, one will cause significantly larger shifts in the inferred model parameters (both astrophysical and HMF).

Finally, following the inclusion of UV LFs into our joint HMF plus astrophysical parameter model, we still find  $\sim 2-3\times$  larger marginalized uncertainties on our astrophysical parameters. However, importantly, we recover all input astrophysical parameters to within our 68th percentile uncertainties. Therefore, by jointly recovering the astrophysical and HMF model parameters we can recover unbiased constraints on the input parameters. Although the relative uncertainties are notably larger when comparing to assuming a fixed HMF, they are conservative and more importantly agnostic to the choice of HMF. Further, we also recovered the marginalized posteriors on the underlying HMF and although they were relatively tightly constrained, they were broader than the variation between the various HMFs considered earlier in this work and thus we are unable to perform model selection to distinguish the input (ST) HMF.

In future, we could expand on this work by utilizing the Bayesian evidence to perform model selection to (i) attempt to rule out certain incorrectly assumed choices of the HMF or (ii) to determine the necessary complexity of a generic HMF beyond the five parameter model considered within this work. Alternatively, rather than jointly sampling the astrophysical and HMF parameters, if we are only interested in the astrophysical parameters we could consider performing Bayesian model averaging of the extracted posteriors obtained from a broad range of HMF models.

## ACKNOWLEDGEMENTS

We would like to thank the anonymous referees whose comments improved this manuscript. Parts of this research were supported by the Australian Research Council Centre of Excellence for All Sky Astrophysics in 3 Dimensions (ASTRO 3D), through project number CE170100013. JM was supported by an appointment to the NASA Postdoctoral Program at the Jet Propulsion Laboratory/California Institute of Technology, administered by Oak Ridge Associated Universities under contract with NASA. YST acknowledges financial support from the Australian Research Council through DECRA Fellowship DE220101520. A.M. acknowledges support from the Ministry of Universities and Research (MUR) through the PRIN project ‘Optimal inference from radio images of the epoch of reionization’ as well as the PNRR project ‘Centro Nazionale di Ricerca in High Performance Computing, Big Data e Quantum Computing’. This work was performed on the OzSTAR national facility at Swinburne University of Technology. The OzSTAR programme receives funding in part from the Astronomy National Collaborative Research Infrastructure Strategy (NCRIS) allocation provided by the Australian Government, and from the Victorian Higher Education State Investment Fund (VHESIF) provided by the Victorian Government.

*Software:* NUMPY (Harris et al. 2020), SCIPY (Virtanen et al. 2020), H5PY (Collette 2013), MATPLOTLIB (Hunter 2007), CORNER (Foreman-Mackey 2016), SEABORN (Waskom et al. 2017).

## DATA AVAILABILITY

The data underlying this article will be shared on reasonable request to the corresponding author.

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**APPENDIX A: ASSESSING NETWORK COVERAGE**

One of the significant advantages of SBI like MNRE is that once the network is trained they provide rapid evaluation of the posteriors for any given input data. That is, we can perform inference for a large number of input models (not used in the training of the original network) sampling our prior volume and determine the frequency with which each model falls within the predicted posteriors determined by our trained network. Ultimately, this enables us to quantify the coverage of the network (e.g. Cole et al. 2022), which is a considerably more robust concept than the typical convergence criteria used by traditional MCMC approaches (e.g. Betancourt 2019; Roy 2020).

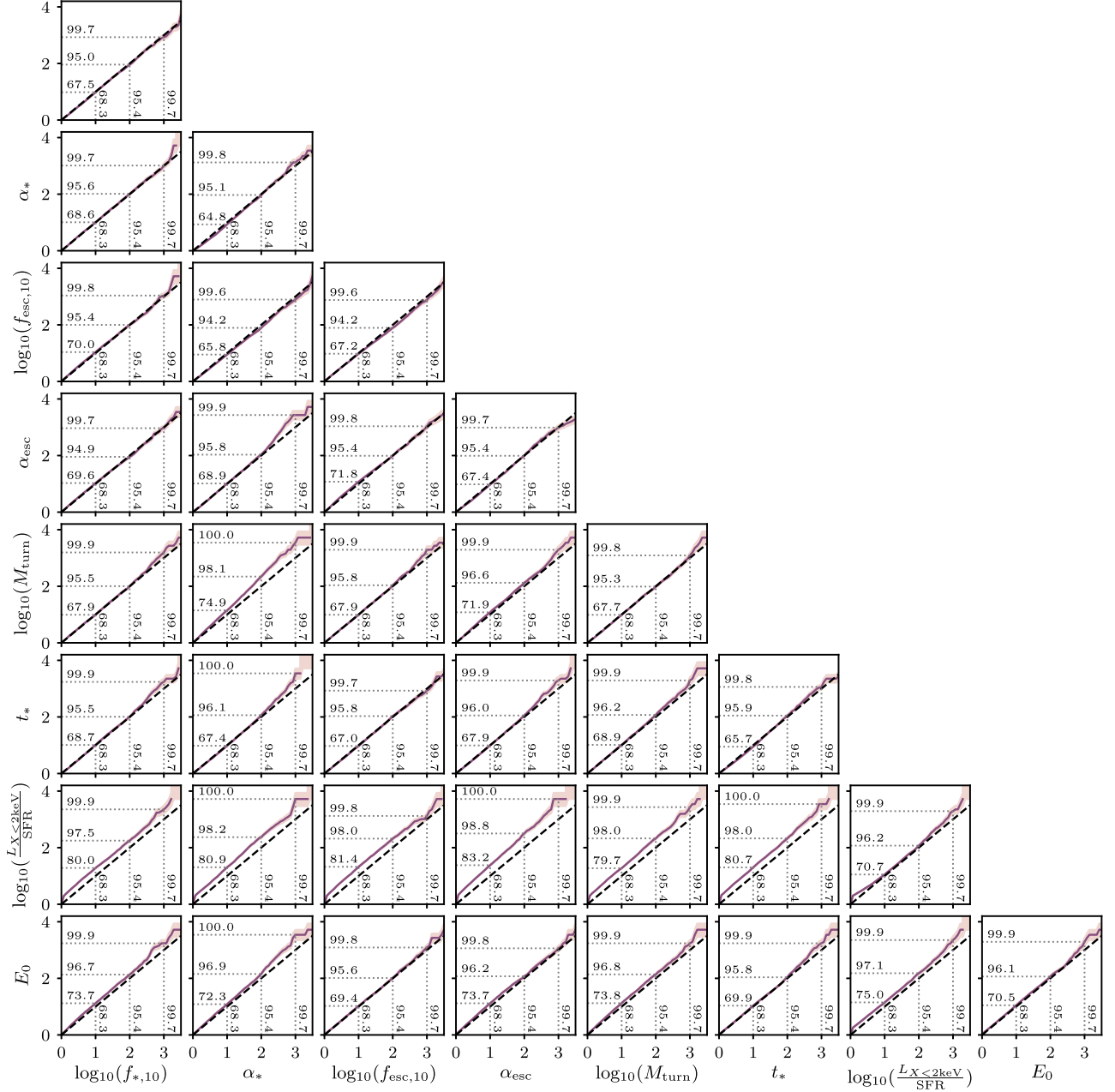
First, we define the function,  $\Theta_{\hat{p}(\hat{\theta}|\mathbf{x}_i)}(1 - \alpha)$ , which gives the  $(1 - \alpha)$  highest probability density region (HPDR) for our estimated posterior,  $\hat{p}(\hat{\theta}|\mathbf{x}_i)$ , given the input mock data,  $\mathbf{x}_i$ , and known input parameters,  $\theta_i^*$ . For example, a 95 percent HPDR would have an error rate of  $\alpha = 0.05$ . Then, for an independent set of  $n$  randomly drawn models,  $(\mathbf{x}_i, \theta_i^*)$ , we determine

$$1 - \hat{\alpha} = \frac{1}{n} \sum_{i=1}^n \mathbb{1}[\theta_i^* \in \Theta_{\hat{p}(\hat{\theta}|\mathbf{x}_i)}(1 - \alpha)], \quad (\text{A1})$$

the actual error rate of the HPDR given our estimated posterior.

The quantities  $\alpha$  ( $\hat{\alpha}$ ) can be re-parametrized in terms of a new variable,  $z$ , which corresponds to the  $1 - \alpha/2$  ( $1 - \hat{\alpha}/2$ ) quantile of the standard normal distribution. In this way, the 1, 2, 3 $\sigma$  regions correspond to  $z = 1, 2, 3$  with  $1 - \alpha = 0.6827, 0.9545, 0.9997$ . The uncertainties on  $\hat{\alpha}$  are determined using the Jeffreys interval (Cole et al. 2022).<sup>12</sup> In Fig. A1, we present the empirical expected coverage probability of our trained network for our mock 21-cm observation assuming an ST HMF as a function of confidence levels for all 1D and 2D marginalized posteriors. For this, we used a total of 5000 independent models (not used in the training of our MNRE network) to ascertain the coverage probabilities. The goal is to achieve a network coverage (purple line) that follows the black dashed curve (perfect coverage). Deviations of the resultant network coverage above the dashed black line implies our network is conservative with respect to the recovered posteriors whereas deviations below are considered overconfident. For our particular set-up, we recover excellent network coverage, with near perfect performance for most 2D parameter combinations.

<sup>12</sup>This interval is the 68.27 percent central interval of a Beta distribution defined by the parameters  $(n - k + 1/2, k + 1/2)$ , where  $n$  is the total number of samples from the joint model and  $k$  is the number of times the HPDR predicted by the network does not contain the true astrophysical parameters.



**Figure A1.** The empirical expected coverage probability of our trained MNRE network with SWYFT (vertical axis) as a function of the confidence level (horizontal axis). The purple line corresponds to the coverage of our network, with perfect coverage denoted by the diagonal black dashed line, while dotted lines denote the 68th, 95th and 99.7th percentiles while the shaded region corresponds to the Jeffrey's interval (see the text for further details).

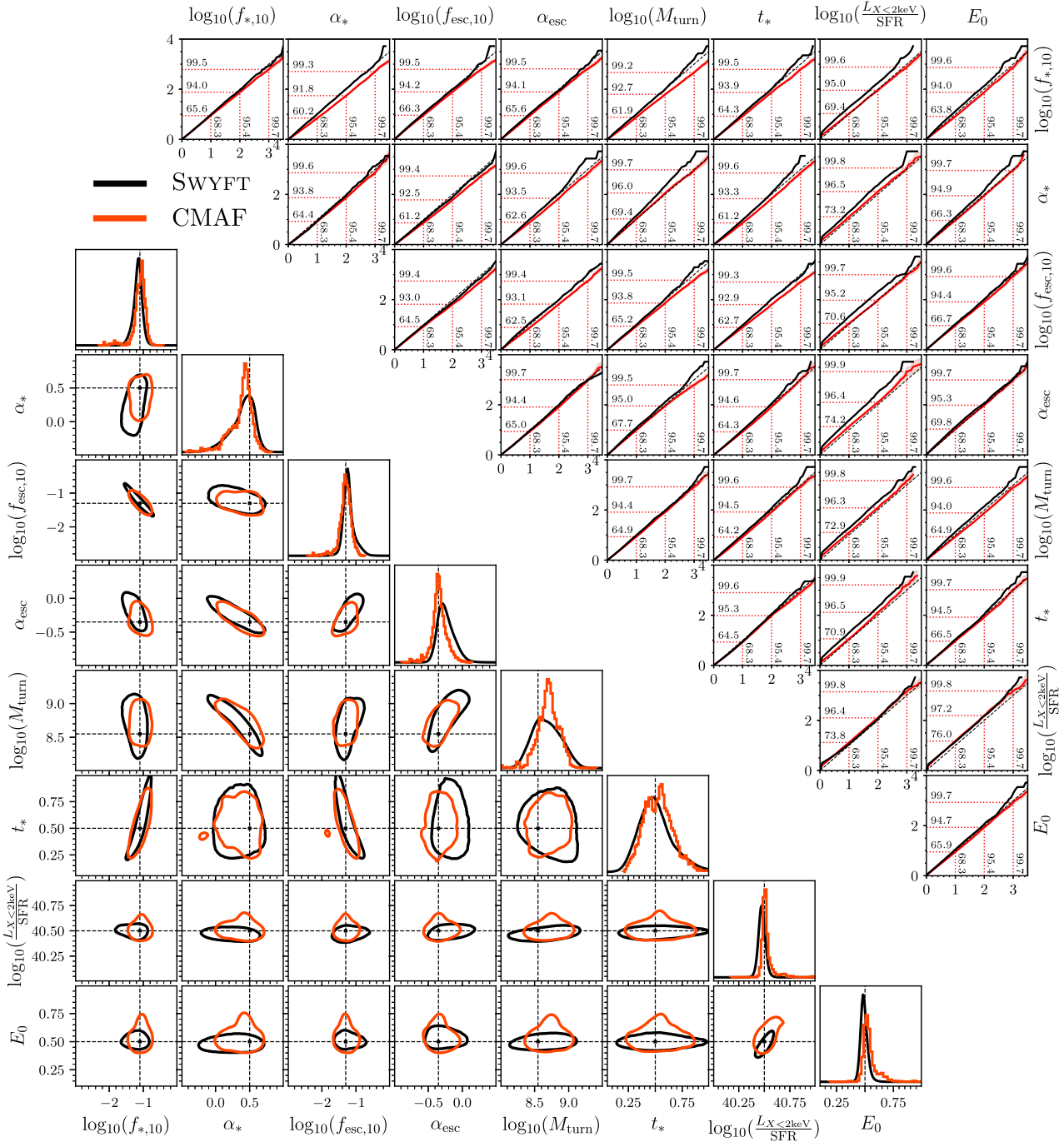
## APPENDIX B: COMPARISON OF SBI APPROACHES

During the completion of this work we simultaneously employed MNRE with SWYFT to learn the likelihood-to-evidence ratio while also applying the conditional masked-autoregressive flow (CMAF) model from Prelogović & Mesinger (2023) to explicitly learn the likelihood. CMAF is a density estimator which performs a series of linear transformations of normal random variables to estimate a target distribution (see e.g. Papamakarios et al. 2021 for a recent

review). Using the publicly available code, 21CMLIKELIHOODS<sup>13</sup> we performed SBI on the same simulated data set of models as outlined in Section 3.2 to learn our likelihood. From this, we then use EMCEE (Foreman-Mackey et al. 2013) assuming the same priors to obtain our posteriors given the same mock 21-cm PS observation.

In the lower left corner of Fig. B1, we provide the recovered 1D and 2D marginalized posteriors from SWYFT (black curves) and after learning the likelihood using CMAF (red curves). Below the diagonal, we provide the 95th percentile joint marginalized posteriors. Generally speaking we recover consistent results between

<sup>13</sup><https://github.com/dprelogo/21cmLikelihoods>



**Figure B1.** Comparison of the two different SBI approaches: SWYFT (black contours and curves) and conditional masked-autoregressive flow (CMAF: red contours and curves). In the lower left corner, we compare the 1D and 2D marginalized posteriors of our fiducial astrophysical model assuming our mock 21-cm PS observation with the SKA. In the upper right corner, we provide the empirical expected coverage probability (see Fig. A1 for more details). Here, the reported 68th, 95th, and 99.7th percentiles coverage probabilities corresponds to the CMAF trained likelihood.

the two different SBI approaches, with SWYFT typically producing slightly broader (more conservative) posteriors. Given the vastly different methodologies employed by both methods to achieve similar results, this yields additional confidence in both approaches. To more rigorously compare the two SBI approaches in the upper right corner of Fig. B1, we also compare the empirical expected coverage probabilities (see Appendix A) of the CMAF network (red

curves) and our previous MNRE network from SWYFT (black curves). Here, the reported coverage probabilities correspond to those from using the CMAF approach. In general, we find SWYFT to be more conservative as indicated by the black curve extending further above the CMAF coverage (red curve). For a few parameter combinations, we find the CMAF approach to be very slightly overconfident (the red curve dips below the black dashed diagonal line; for example along

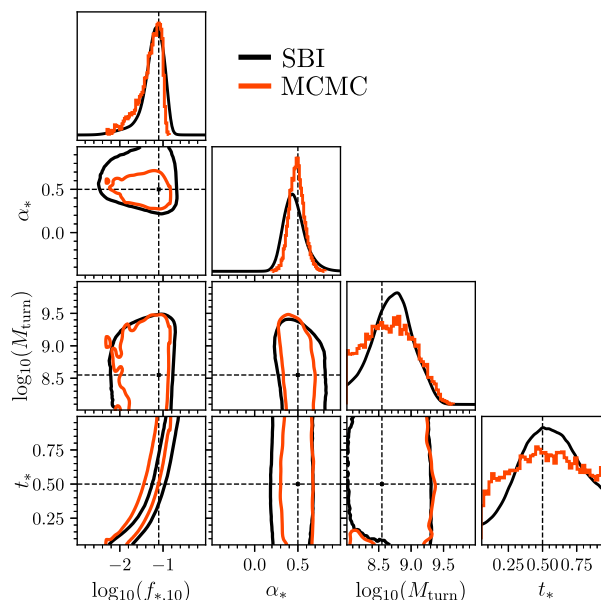
the top row) which is consistent with the narrower (tighter) posteriors shown in the lower left corner of Fig. B1. In the end, we found SWYFT to be more computationally efficient, conservative and also more stable to train and thus adopted it for the remainder of this work. It is likely that the CMAF approach might require additional tweaks to the available hyperparameters or additional ensemble learning or active learning as explored in (Alsing et al. 2019) to boost its stability.

### APPENDIX C: MODIFYING UV LFS FOR SBI

SBI requires a stochastic simulator of the modelled data given a set of input parameters. However, in 21CMFAST our UV LFs are derived from analytic functions as outlined in equation (6). Therefore, to translate the observational uncertainties from real-world observations into our simulated UV LFs we simply add random noise to each  $M_{UV}$  bin (see Section 4.3). Here, we verify this approximation by using SWYFT to obtain our marginalized posteriors on our astrophysical model parameters using only the UV LF data in our training set and compare the results to an MCMC where we produce the UV LFs from 21CMFAST on-the-fly using EMCEE and assume a simple analytic likelihood (i.e.  $\chi^2$ , Park et al. 2019).

In Fig. C1, we compare the 1D and 2D marginalized posteriors (95th percentiles) using SWYFT (black curves) and our direct MCMC approach (red curves). The consistency of the posteriors following the two approaches provides confidence that our procedure for adding noise to the analytic UV LFs in order to be able to include them within our SBI framework is sufficiently accurate. Note, we do observe slightly tighter 1D marginalized posteriors for  $M_{turn}$  and  $t_*$  with SWYFT implying improved constraining power over direct MCMC, however, the differences are fairly minor. Nevertheless, these differences likely arise because of the implicit likelihood that the MNRE approach of SWYFT is extracting, which demonstrates the importance of going beyond the simple  $\chi^2$  likelihood form of traditional MCMCs. The differences between the likelihoods are

non-trivial, but in applying MNRE we always reduce corresponding biases resulting in more precise posteriors.



**Figure C1.** Comparison of our marginalized posteriors obtained from only considering UV LFs. The black curves correspond to performing SBI with SWYFT after modifying the simulated data to directly include observational uncertainties and the red curves correspond to a direct MCMC assuming an analytic form for the likelihood.

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